

## On Minimax Invariant Estimates of the Transformation Parameter

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Among decision problems in the statistical theory, there are many cases where the distribution space  $\Omega$  remains invariant under any transformation belonging to a group  $G$  of transformations defined on the sample space. For example, if  $(x_1, x_2, \dots, x_n)$  is the sample of a fixed size from a normal distribution with unknown mean and unknown variance, then the normality of the distribution of the sample remains invariant under any transformation  $(x_1, x_2, \dots, x_n) \rightarrow (ax_1 + b, ax_2 + b, \dots, ax_n + b)$ . Of course, the parameters involved in the distribution function vary within  $\Omega$  according as the transformation does. If an initial distribution is given in each type, the parameters can be considered as induced from the initial distribution under transformations on the sample space. Taking account of this fact, such a set of parameters should be called a transformation parameter. The pair of the location and scale parameters is an example of such parameters.

The group theoretic properties of transformations on the sample space were often used as a key of solving some special problems (*e.g.* Pitman [18, 19], Wolfowitz [23] and the author [13, 14]). In the general problem of estimation of the transformation parameter, are these properties powerful tools?

On the other hand, if the observations are our only source of information of the value of the transformation parameter, a satisfactory solution of the problem should remain invariant under any transformation of the group (the principle of invariance). Hence it seems desirable that the minimax solution of such a problem (if exists) is also invariant. Does it hold true?

The third question is: Can the invariant measure on  $\Omega$  or a sequence of its truncated measures be used as a least favourable a priori distribution, relative to which the minimax invariant decision function is a Bayes solution? If the answer of this question is affirmative, the classical Bayes' method of solving the problem by using the uniform distribution as an a priori distribution becomes to have a reasonable foundation from the viewpoint of the subjective probability (see de Finetti [7]). In this connection Karlin [12] has proved a theorem on the theory of games that the invariant Haar measure on a compact group as a space of pure strategies is a minimax mixed strategy.

The above three questions are answered, in this paper, in the affirmative under certain conditions.

A curious one of these conditions is that the group  $\bar{G}$  of transformations on  $\Omega$  induced by  $G$  is an  $A$ -group (for definition, see page 45). But this condition seems to be unavoidable, since the special groups under which Hunt and Stein have proved the existence of the most stringent test are all  $A$ -groups (see Lehmann [16] and our Theorem 2.1)

The existence of the invariant measure on  $\Omega$  is guaranteed by Mibu's theorem, as an extension of Haar measure on a locally compact group. Mibu [17] has developed his theory of the invariant measure on a locally compact and  $\sigma$ -compact uniform space admitting a group of transformations, and introduced a topology of the group such that the given space becomes a coset space. His theory gives us an important tool of our discussions. Hereupon we need to define the topology of  $\Omega$ . Fortunately the absolute variation  $V(P, P')$  of the difference of two distributions  $P$  and  $P'$  of  $\Omega$  can be properly used as a metric of  $\Omega$ , because  $V(P, P')$  becomes larger according as the discrimination of  $P$  against  $P'$  by testing procedure is more difficult (see Kudō [13, 15]).

In the problem of estimation of the transformation parameter, it is not a considerable restriction to suppose that the decision space admits a group  $\tilde{G}$  isomorphic or homomorphic to  $\bar{G}$ , and that the loss function  $W(P, \alpha)$  remains invariant under any simultaneous and corresponding transformation  $\sigma \in G$  of the distribution  $P$  and of the decision  $\alpha$ , *i. e.*  $W(\bar{\sigma}P, \tilde{\sigma}\alpha) = W(P, \alpha)$  for any  $\sigma \in G$ . For instance, when the decision space coincides with  $\Omega$ ,  $V$  is one of such functions, and hence every function on  $\Omega \times \Omega$  only through  $V$  can be also used as a loss function. In this paper we restrict ourselves to consider the problem of estimation of the transformation parameter under a loss function mentioned above. However, for the generality of discussions, we speak of decision functions instead of estimates. By doing so, our results stand more close to Lehmann's theorem [16, Th. 9.1].

The concept of the invariant estimate was introduced by Fisher [8] at an early stage in the development of statistical theory, and has proved extremely fruitful. In 1938 Pitman [18] proposed a general method of constructing the invariant estimate of the location and scale parameters (simultaneously and separately) by using the notion of the "fiducial function". (See example *c*). The rearrangement of Pitman's estimate from the standpoint of the theory of decision functions was given by Girshick and Savage [9]. They have proved that Pitman's estimate is minimax with respect to the quadratic loss function  $(\theta - a)^2$ , where  $\theta$  is the parameter and  $a$  is the estimated value. In Blackwell-Girshick's 1954 Book [4], the existence of the minimax invariant es-

timate of the location parameter is proved in the case of discrete variables in Euclidean space under a certain general condition of the loss function  $W(\theta - a)$  which is satisfied either by every bounded  $W(u)$  or by every continuous  $W(u)$  having the infinite limits at  $u = \pm \infty$ . (See Example *b*). Discussions of the abstract transformation parameter problem have been already developed in 1950 by Hunt, Stein and Lehmann (see Lehmann [16]), but they concerned only the problem of testing hypotheses.<sup>(0)</sup>

In the first section of this paper, we refer to Mibu's result without any proof and give the consideration about invariant statistics, which is inevitable in the following studies. Section 2 is the trunk of the paper. We shall give there a straightforward extension (Theorem 2.2) of Blackwell-Girshick's result [4, Theorem 11.3.1], which answers the above three questions. The problem of estimation of cosets, like that of estimation of the mean from the observation of normal variables with unknown variance and a problem of estimation by interval are discussed in the last of the section (Theorems 2.6, 2.7). It happens often that the sample space admitting a group  $G$  has an *a priori* topological structure. Section 3 is devoted to the proof of the homomorphism between two topological groups  $G$  and  $\bar{G}$ , when  $P$  is absolutely continuous with respect to the invariant measure on the sample space. The results obtained in Section 3 prove their worth in Section 4, in which some illustrative examples are gathered. The last example *f*) of Section 4 is given as an open problem.

## 1. Group of transformations of the sample space.

### 1.1. Topology of the parameter group. Mibu's Theorem.

Let  $(X, \mathfrak{B}, P)$  be a probability space,<sup>(1)</sup> and  $G$  be a group of 1:1 measurable transformations<sup>(2)</sup>  $e, \sigma, \tau, \rho, \dots$  of  $X$  onto itself ( $e$  is the invariant transformation). Consider the probability measure  $P(\sigma^{-1}B)$  on the measurable space  $(X, \mathfrak{B})$ , and denote it by  $\bar{\sigma}P(B)$ . For such measures  $\bar{\sigma}P$ , it holds that  $\bar{\sigma}(\bar{\rho}P) = (\bar{\sigma\rho})P$ , and we may write as  $\bar{\sigma\rho}P$  instead of  $(\bar{\sigma\rho})P$ . Since

$$G_0 = \{\sigma : \bar{\sigma\rho}P = \bar{\rho}P \text{ for every } \rho \in G\}$$

is a normal subgroup of  $G$ , the factor group  $\bar{G} = G/G_0$  may be considered as a group of transformations of the space  $\Omega = \{\bar{\sigma}P : \sigma \in G\}$ . From this

(0) [24] reports that Peisakoff obtained some results concerning the invariant minimax decision function. They are perhaps very closely related to our task. I regret that they are unpublished.

(1) We assume that every  $\sigma$ -field of subsets of the spaces appearing in this paper contains every single point set, and is generated by a countable number of subsets of the space.

(2) If  $B \in \mathfrak{B}$  implies  $\sigma^{-1}B \in \mathfrak{B}$ , then we call the transformation  $\sigma$  to be measurable.

fact, we shall regard  $\bar{\sigma}, \bar{\tau}, \dots$  operating on elements of  $\Omega$  as an element of the group  $\bar{G}$ .

For the subject of this paper, it is convenient to use our terminologies as follows:

$\Omega$  = the distribution space,

$X$  = the sample space,

$\bar{G}$  = the parameter group,

an element of  $\Omega$  = a distribution,

an element of  $X$  = a sample point,

an element of  $\bar{G}$  = a transformation parameter.

As we have seen in our previous papers [13, 15], the absolute variation

$$V(\bar{\sigma}P, \bar{\tau}P) = \sup_{B \in \mathfrak{B}} [\bar{\sigma}P(B) - \bar{\tau}P(B)] - \inf_{B \in \mathfrak{B}} [\bar{\sigma}P(B) - \bar{\tau}P(B)]$$

of the difference of two probability measures  $\bar{\sigma}P$  and  $\bar{\tau}P$  can be used as a "measure" of information of the experiment<sup>(3)</sup>  $(\bar{\sigma}P, \bar{\tau}P)$ . Hence it seems natural for the statistical discussion to use  $V(\bar{\sigma}P, \bar{\tau}P)$  as the metric in  $\Omega$ <sup>(4)</sup>. Thus  $\Omega$  is regarded as a topological space with the metric  $V$  in the following studies.

**Assumption A.** *The metric space  $\Omega$  is locally compact and  $\sigma$ -compact.<sup>(5)</sup> Hence  $\Omega$  is separable and complete.*

**Remark 1.1.** Owing to A. Berger [1], the condition of the separability of  $\Omega$  is equivalent to that  $\Omega$  is dominated by a  $\sigma$ -finite measure  $\mathfrak{l}$ :

$$\sigma P(B) = \int_B p(x : \sigma) \mathfrak{l}(dx), \quad B \in \mathfrak{B}.$$

The general theory of the topology of a group of transformations and of the invariant measure on a uniform space was developed by Yoshimichi Mibu [17]. However we shall cite here his results in a more restricted form.

**Theorem 1.1.** (Mibu) *Suppose that  $\Omega$  is a metric space<sup>(6)</sup> satisfying Assumption A, and that  $\bar{G}$  is a group of isometric transformations of  $\Omega$  onto itself, with respect to which  $\Omega$  is homogeneous.<sup>(7)</sup> Taking as a neighbourhood of the neutral element  $\bar{e}$  of  $\bar{G}$  the subsets*

(3) See Blackwell [3] and Kudō [15].

(4) We can see easily that  $V(\sigma P, \tau P)$  is a  $\bar{G}$ -invariant metric: (1)  $V(\bar{\sigma}P, \bar{\tau}P) = 0$  with the equality when and only when  $\bar{\sigma}P = \bar{\tau}P$ , (2)  $V(\sigma P, \bar{\tau}P) = V(\bar{\tau}P, \sigma P)$ , (3)  $V(\sigma P, \bar{\tau}P) \leq V(\sigma P, \bar{\rho}P) + V(\bar{\rho}P, \bar{\tau}P)$  and (4)  $V(\bar{\rho}\bar{\tau}P, \bar{\rho}\sigma P) = V(\bar{\tau}P, \bar{\sigma}P)$  for all  $\bar{\rho} \in \bar{G}$ .

(5) A topological space  $E$  is said to be  $\sigma$ -compact, if  $E$  is the sum of a countable number of compact subsets.

(6) It does not assumed in Theorems 1.1-1.4 that  $\Omega$  is a distribution space. These theorems hold when  $\Omega$  is an arbitrary metric space with a  $\bar{G}$ -invariant metric  $V$ .

(7) A space  $E$  is said to be homogeneous with respect to a group  $G$  of transformations of  $E$  if, given any two points of the space, there is an element of the group which transforms one into the other.

$$U(K, \epsilon) = \{ \bar{\sigma} : V(\bar{\sigma}\bar{\rho}P, \bar{\rho}P) < \epsilon \text{ for every } \bar{\rho}P \in K \},$$

where  $K$  is a compact subset of  $\Omega$  and  $\epsilon$  is a real positive number,  $\bar{G}$  becomes a locally totally bounded topological group satisfying the first axiom of countability<sup>(8)</sup> and the mapping  $(\bar{\rho}P, \bar{\sigma}) \rightarrow \bar{\sigma}^{-1}\bar{\rho}P$  of  $\Omega \times \bar{G}$  onto  $\Omega$  is continuous.

A necessary and sufficient condition of  $\bar{G}$  to be totally bounded is that  $\Omega$  is compact.

**Assumption B.**<sup>(9)</sup> If  $\bar{\sigma}_1\bar{\rho}P, \bar{\sigma}_2\bar{\rho}P, \dots$  is a Cauchy sequence of points  $\in \Omega$  for every  $\bar{\rho} \in \bar{G}$ , then there exists a  $\bar{\sigma}_0 (\in \bar{G})$  such that  $\bar{\sigma}_n\bar{\rho}P$  tends to  $\bar{\sigma}_0\bar{\rho}P$  in the sense of the metric  $V$  as  $n \rightarrow \infty$  for every  $\bar{\rho} \in \bar{G}$ .

**Theorem 1.2.** (Mibu). Suppose, in addition to the hypotheses of Theorem 1.1, that  $\bar{G}$  fulfills Assumption B. Then  $\bar{G}$  is complete, locally compact and  $\sigma$ -compact. The compactness of  $\Omega$  is a necessary and sufficient condition of the compactness of  $\bar{G}$ . Since  $\bar{G}$  is metrisable [5, Ch. IX, §3], the separability of  $\bar{G}$  follows from the  $\sigma$ -compactness of  $\bar{G}$ .

The subgroup  $\bar{H}$  consisting of the transformations under which  $P$  remains invariant is closed in the topological group  $\bar{G}$  under Assumption A. In fact, let  $\bar{\sigma} \notin \bar{H}$ . Since  $\epsilon = V(\bar{\sigma}P, P) > 0$ , the neighbourhood  $U(\{\bar{\sigma}P\}, \epsilon/2)$  of  $\bar{\sigma}$  have no common element with  $\bar{H}$ , that is,  $\bar{H}$  is closed in  $\bar{G}$ .

Thus we can treat the space  $\bar{G}/\bar{H}$  of the left cosets  $\bar{\sigma}\bar{H}$  as a topological space, in which the neighbourhoods of  $\bar{\sigma}\bar{H}$  have a form  $U(K, \epsilon)\bar{\sigma}\bar{H}$ . It is evident that the mapping  $\alpha : \bar{\sigma}\bar{H} \rightarrow \bar{\sigma}P$  of  $\bar{G}/\bar{H}$  onto  $\Omega$  is 1:1, and moreover we have the following

**Theorem 1.3.** The mapping  $\alpha$  is topological; in the other words,  $\Omega$  is homeomorphic to  $\bar{G}/\bar{H}$  under the mapping  $\alpha$ .

(8) Since  $\bar{\tau}U(\{P\}, \epsilon/3)P = \{\bar{\rho}P : V(\bar{\rho}P, \bar{\tau}P) < \epsilon/3\}$ , and since  $K$  is compact, there is a finite number of elements  $\bar{\tau}_1, \bar{\tau}_2, \dots, \bar{\tau}_n$  of  $\bar{G}$  such that

$$K \subset \bigcup_{i=1}^n \bar{\tau}_i U(\{P\}, \epsilon/3)P.$$

Hence we have

$$\begin{aligned} \bigcap_{i=1}^n U(\{\bar{\tau}_i P\}, \epsilon/3) &= \bigcap_{i=1}^n \bar{\tau}_i U(\{P\}, \epsilon/3)\bar{\tau}_i^{-1} \\ &\subset \bigcap_{i=1}^n \bigcap_{\sigma \in U(\{P\}, \epsilon/3)} \bar{\tau}_i \bar{\sigma} U(\{P\}, \epsilon) \bar{\sigma}^{-1} \bar{\tau}_i^{-1} \\ &= \{ \bar{\rho} : V(\bar{\sigma}\bar{\tau}_i P, \bar{\tau}_i \bar{\sigma} P) < \epsilon \text{ for } \bar{\sigma} \in U(\{P\}, \epsilon/3), i=1, 2, \dots, n \} \\ &\subset \{ \bar{\rho} : V(\bar{\rho}\bar{\tau}_i P, \bar{\tau}_i P) < \epsilon \text{ for every } \bar{\tau}_i P \in K \} \\ &= U(K, \epsilon). \end{aligned}$$

This shows that  $\{U(\{\bar{\tau}P\}, r) : \bar{\tau} \in \bar{G}, r = \text{a positive number}\}$  forms a complete system of neighbourhoods of the neutral element  $\bar{e}$  in  $\bar{G}$ .

(9) If  $\bar{G}$  is commutative or if  $G$  is the group of all isometric transformations on  $\Omega$ , then this Assumption is superfluous.

**Proof.** It is evident that  $\alpha$  is continuous, since the mapping  $\bar{\sigma} \rightarrow \bar{\sigma}P$  is continuous (the definition of the topology of  $\bar{G}$ ), and since the mapping  $\bar{\sigma} \rightarrow \bar{\sigma}\bar{H}$  is open.

That  $\alpha$  is open follows from the fact that

$$\alpha[U(\bar{\sigma}P, \varepsilon)\bar{\sigma}\bar{H}] = \{\bar{\tau}\bar{\sigma}P : V(\bar{\tau}\bar{\sigma}P, \bar{\sigma}P) > \varepsilon\}.$$

**Corollary.** If every transformation of  $\bar{G}$ , except for the neutral element, transforms  $P$  to the other point of  $\Omega$ , then  $\Omega$  is homeomorphic to  $\bar{G}$ .

**Theorem 1.4.** (Mibu). *Suppose that  $\Omega$  and  $\bar{G}$  satisfy the hypotheses of Theorem 1.2. There exists an outer measure  $\xi^*$  on  $\Omega$ , which satisfies the following conditions:*

- i)  $\xi^*(\bar{\sigma}A) = \xi^*(A)$ ,  $\bar{\sigma} \in \bar{G}$ ,  $A \subset \Omega$ .
- ii) *Every Borel set is measurable.*<sup>(10)</sup>
- iii) *For any subset  $A$  of  $\Omega$ , there exists a Borel set  $B$  such that*

$$\xi^*(A) = \xi(B).$$

(we shall write as  $\xi$  the outer measure  $\xi^*$  of measurable set).

- iv) *If  $A$  is compact, then  $\xi(A) < \infty$ .*
- v)  *$\xi^*(\Omega) < \infty$  if and only if  $\Omega$  is compact.*
- vi)  *$\xi$  is a continuous measure; i. e. there corresponds a neighbourhood  $U$  of  $\bar{e}$  of  $\bar{G}$  for every positive  $\varepsilon$  and every measurable subset  $A$  ( $\xi(A) < \infty$ ) of  $\Omega$ , such that*

$$\xi[(\bar{\sigma}A \cup A) - (\bar{\sigma}A \cap A)] < \varepsilon, \bar{\sigma} \in U,$$

vii) *If  $\mu^*$  is a left invariant Haar measure on  $\bar{G}$  and  $f(\bar{\rho}P)$  is a  $\xi^*$ -measurable function on  $\Omega$ , then  $f(\bar{\sigma}^{-1}\bar{\rho}P)$  is a  $\xi^*\mu^*$ -measurable function on  $\Omega \times \bar{G}$ .*

viii) *The outer measure satisfying the conditions i), ii) and iii) is unique.*

ix) *It is a necessary and sufficient condition for a measure  $\pi$ , defined on the class of the Borel subsets of  $\Omega$ , being absolutely continuous with respect to  $\xi$  that there corresponds a neighbourhood  $U$  of  $\bar{e}$  of  $\bar{G}$  for every  $\varepsilon > 0$  and every compact subset  $K$  of  $\Omega$ , such that*

$$[\text{absolute variation of } (\bar{\sigma}\pi - \pi) \text{ on } K] < \varepsilon, \bar{\sigma} \in U,$$

where  $\bar{\sigma}\pi(A) = \pi(\bar{\sigma}^{-1}A)$  for every Borel set  $A \subset \Omega$ .

x) *For two measurable subsets  $A$  and  $B$  of positive  $\xi$ -measure, there is an element  $\bar{\tau} \in \bar{G}$  such that  $\xi(\bar{\tau}A \cap B) > 0$ .*

(10) Under Assumption A, the concept of Borel sets is equivalent to that of Baire sets. See P. R. Halmos [10; p. 220].

This measure  $\xi$  is regarded as an invariant measure on  $\overline{G}/\overline{H}$ .

Since the closed subgroup  $\overline{H} = \{\overline{\sigma} : \overline{\sigma}P = P\}$  of the locally compact group  $\overline{G}$  is also locally compact, there exists a left invariant Haar measure  $\nu$  on  $\overline{H}$ . For any Borel measurable function  $f(\overline{\sigma})$  on  $\overline{G}$ , the integral

$$(1.1) \quad \int_{\overline{G}/\overline{H}} \xi(d\overline{\sigma H}) \int_{\overline{H}} f(\overline{\sigma\tau}) \nu(d\tau)$$

is invariant under any transformation  $\overline{\sigma} \rightarrow \overline{\rho\sigma}$ . In fact, writing

$$f^*(\overline{\sigma H}) = \int_{\overline{H}} f(\overline{\sigma\tau}) \nu(d\tau),$$

we have

$$(1.2) \quad \int_{\overline{G}/\overline{H}} f^*(\overline{\rho\sigma H}) \xi(d\overline{\sigma H}) = \int_{\overline{G}/\overline{H}} f^*(\overline{\sigma H}) \xi(d\overline{\sigma H}).$$

(1.2) shows the  $\overline{G}$ -invariance of the integral (1.1). Therefore, by the uniqueness of the left invariant Haar measure on  $\overline{G}$ , we have

$$\int_{\overline{G}} f(\overline{\sigma}) \mu(d\overline{\sigma}) = \int_{\overline{G}/\overline{H}} \xi(d\overline{\sigma H}) \int_{\overline{H}} f(\overline{\sigma\tau}) \nu(d\tau),$$

that is, we may write without any confusion that

$$\int_{\overline{G}} f(\overline{\sigma}) \mu(d\overline{\sigma}) = \int_{\Omega} \xi(d\overline{\sigma P}) \int_{\overline{H}} f(\overline{\sigma\tau}) \nu(d\tau).$$

We shall denote by  $\mathfrak{B}$  the  $\sigma$ -field of all Borel subsets of the parameter group  $\overline{G}$  throughout the remainder of this paper.

### 1.2. Invariant statistics.

Every measurable mapping  $t(x)$  of the sample space  $(X, \mathfrak{B})$  into another measurable space  $(T, \mathfrak{T})$  induces a decomposition of  $X$  into disjoint measurable sets  $X_t = \{x : t(x) = t\}$ ,  $t \in T$ . (See footnote (1)). Writing

$$(\overline{\sigma}P)_r(E) = \overline{\sigma}P(t^{-1}(E)) \text{ for } E \in \mathfrak{T},$$

$\overline{\sigma}P$  is resolved as

$$\overline{\sigma}P(B \cap t^{-1}(E)) = \int_E \overline{\sigma}P(B : t) (\overline{\sigma}P)_r(dt)$$

for  $B \in \mathfrak{B}$  and  $E \in \mathfrak{T}$ , where  $\overline{\sigma}P(B : t)$  is a non-negative measurable function on  $T$  depending on  $B \in \mathfrak{B}$  and on  $\overline{\sigma}P$ . In general,  $\overline{\sigma}P(B : t)$  is not uniquely determined. However if we can choose  $\overline{\sigma}P(B : t)$  such that  $\overline{\sigma}P(B : t)$  coincides with  $\overline{\sigma}P(B \cap X_t : t)$  for  $\nu_r$ -almost all  $t \in T$  and for all  $\overline{\sigma}P \in \Omega$ , and if  $\overline{\sigma}P(B \cap X_t : t)$  is a probability measure on  $(X_t, \mathfrak{B}_t)$ ,  $\mathfrak{B}_t = \{B \cap X_t : B \in \mathfrak{B}\}$ , then we call  $t(x)$  a statistic and say that  $\overline{\sigma}P(B : t)$  is a

conditional probability of  $B$  given that  $t(x)=t$ . The set of all  $(\bar{\sigma}P)_T$  is denoted by  $\Omega^T$ .

If the ranges of two statistics  $t(x)$  and  $t'(x)$  are in a 1:1 correspondence  $\beta$  as measurable spaces such that the set  $\{x:\beta(t(x))\neq t'(x)\}$  is of  $\mathbb{I}$ -measure zero, then we say that  $t(x)$  is equivalent to  $t'(x)$ .

Let  $g$  be a subgroup of  $G$ , and  $T$  be the class of the subsets  $gx = \{x' : x' = \sigma x, \sigma \in g\}$  of  $X$ . If the class  $\mathfrak{X}$  of the measurable subsets of  $T$  can be defined such that the mapping  $t : x \rightarrow gx$  is a statistic, then a statistic, equivalent to  $t(x)$ , is said to be  $g$ -invariant, or, i.e.,  $t(x)$  is  $g$ -invariant if and only if

$$x' = \sigma x \text{ and } \sigma \in g \text{ imply } t(x') = t(x) \text{ } \mathbb{I}\text{-almost everywhere on } X.$$

Hence we have, for a  $g$ -invariant statistic  $t(x)$ ,

$$(\bar{\tau}\bar{\sigma}P)_T = (\bar{\sigma}P)_T, \quad \tau \in g \text{ and } \sigma \in G,$$

and

$$\bar{\tau}\bar{\sigma}P(B:t) = \bar{\sigma}P(\tau^{-1}B:t) \quad (\bar{\sigma}P)_T\text{-almost everywhere on } T$$

for all  $\tau \in g$ .

When  $t(x)$  is a statistic invariant under a normal subgroup  $g$  of  $G$ , we can consider a group  $G^T$  of transformations  $\sigma_T$  on  $T$  defined as follows:

$$t(\sigma gx) = t(g\sigma x) = \sigma_T t(x) \text{ } \mathbb{I}\text{-almost everywhere on } X,$$

that is, if  $\sigma \in G$  then

$$\sigma_T t(x) = t(\sigma x) \text{ } \mathbb{I}\text{-almost everywhere on } X.$$

And we have

$$(1.3) \quad (\bar{\sigma}P)_T = \bar{\sigma}_T P_T, \quad \sigma \in G;$$

$$(1.4) \quad \bar{\sigma}P(B \cap t^{-1}(E)) = \int_E \bar{\sigma}P(B:t) \bar{\sigma}_T P_T(dt), \quad \sigma \in G,$$

when  $\bar{\sigma}_T P_T(E) = P_T(\sigma_T^{-1}E)$  for every  $E \in \mathfrak{X}$ . Hence it holds evidently that

$$(1.5) \quad \bar{\sigma}P(B:t) = P(\sigma^{-1} B : \sigma_T^{-1}t) \quad \sigma_T P_T\text{-almost everywhere on } T.$$

Denote

$$G_0^T = \{\sigma : \bar{\sigma}_T \bar{\rho}_T P_T = \bar{\rho}_T P_T \text{ for every } \rho \in G\},$$

and

$$\bar{G}_0^T = \{\bar{\sigma} : \bar{\sigma}_T \bar{\rho}_T P_T = \bar{\rho}_T P_T \text{ for every } \rho \in G\},$$

Since the mapping  $\sigma \rightarrow \bar{\sigma}$  of  $G$  onto  $\bar{G}$  is an algebraic homomorphism with a kernel  $G_0$ , and since  $G_0^T$  is a normal subgroup of  $G$ , the factor group  $\bar{G}^T = G/G_0^T$  is algebraically isomorphic to  $\bar{G}/\bar{G}_0^T$ . On the other hand, since  $\bar{G}^T$  can be regarded as a group of transformations of  $\Omega^T$ , the absolute variation metric  $V_T$  on  $\Omega^T$  reduces  $\bar{G}^T$  to a topological group by Mibu's method stated in Theorem 1.1.



**Theorem 1.5.** *If both of  $\Omega$  and  $\Omega^x$  satisfy Assumptions A and B, then the algebraic homomorphism  $\gamma: \bar{\sigma} \rightarrow \bar{\sigma}_T$  of  $\bar{G}$  onto  $\bar{G}^x$  is open and continuous.*

**Proof.** Let  $F^x$  be a compact subset of  $\Omega^x$ , and  $(\tau_1)_T P_T, (\tau_2)_T P_T, \dots, (\tau_n)_T P_T$  be an  $\varepsilon/3$ -net in  $F^x$ . Choosing one point  $\bar{\tau}_i P$  from each complete inverse image of  $(\tau_i)_T P_T$  under the mapping  $\bar{\tau} P \rightarrow \bar{\tau}_T P_T$  of  $\Omega$  onto  $\Omega^x$ ,  $i=1, 2, \dots, n$ , and denoting  $K = \{\bar{\tau}_1 P, \bar{\tau}_2 P, \dots, \bar{\tau}_n P\}$ , every element  $\bar{\sigma}$  belonging to  $U(K, \varepsilon/3)$  satisfies

$$\begin{aligned} &V_T(\overline{\sigma_T \tau_T P_T}, \bar{\tau}_T P_T) \\ &\leq V_T(\overline{\sigma_T \tau_T P_T}, \overline{\sigma_T (\tau_i)_T P_T}) + V_T(\overline{\sigma_T (\tau_i)_T P_T}, \overline{(\tau_i)_T P_T}) \\ &\quad + V_T(\overline{(\tau_i)_T P_T}, \bar{\tau}_T P_T) \\ &\leq 2V_T(\bar{\tau}_T P_T, \overline{(\tau_i)_T P_T}) + V(\bar{\sigma} \bar{\tau}_i P, \bar{\tau}_i P) \\ &< \frac{2}{3}\varepsilon + \frac{1}{3}\varepsilon = \varepsilon \end{aligned}$$

for every  $\bar{\tau}_T P_T \in F^x$ . Hence  $\bar{\sigma}_T$  belongs to the neighbourhood  $U(F^x, \varepsilon)$  of the neutral element  $\bar{\sigma}_T$  of  $\bar{G}^x$ . Thus the continuity of  $\gamma$  is shown. It follows from Assumptions A and B that  $\gamma$  is an open mapping. (See Pontrjagin [20, Theorem 13]).

By Theorem 1.5 the kernel  $\bar{G}_0^x$  of the mapping  $\gamma$  is closed, and the natural topology of  $\bar{G}/\bar{G}_0^x$  is equivalent to the topology of  $\bar{G}^x = G/G_0^x$ , induced by the metric  $V_T$ .

**Theorem 1.6** *Under the same hypotheses as Theorem 1.5,  $\bar{G}^x$  is isomorphic to  $\bar{G}/\bar{G}_0^x$  under the mapping  $\sigma G_0^x \rightarrow \bar{\sigma} \bar{G}_0^x$  as a topological group.*

In the following we shall give two examples of invariant statistics, each of which will play an important rôle in the later discussions.

a) *A G-invariant statistic  $z(x)$ .* Let  $(Z, \mathfrak{D})$  be a measurable space of the range of  $z(x)$ . Since  $G\sigma = G$  for every  $\sigma \in G$ , every point  $z$  of  $Z$  remains invariant under any  $\sigma_z$ :

$$\begin{aligned} \sigma_z z(x) &= z(\sigma x) = z(x) \\ &\text{for every } \sigma \in G \text{ and } l\text{-almost everywhere on } X. \end{aligned}$$

Hence, from (1.3)–(1.5), we have

$$(1.6) \quad \bar{\sigma}_z P_z = P_z,$$

$$(1.7) \quad \bar{\sigma} P(B \cap z^{-1}(D)) = \int_D \bar{\sigma} P(B : z) P_z(dz), \quad D \in \mathfrak{D}.$$

and

$$(1.8) \quad \bar{\sigma} P(B : z) = P(\sigma^{-1} B : z) \quad P_z\text{-almost everywhere on } Z.$$

From (1.6),  $\Omega^z$  consists only of one point  $P_z$ .

b) A  $G_0$ -invariant statistic  $y(x)$ . Denote by  $(Y, \mathfrak{C})$  the measurable space of the range of  $y(x)$ , and let  $z(y)$  be a  $G$ -invariant statistic on  $(Y, \mathfrak{C})$  onto a measurable space  $(Z, \mathfrak{D})$ . We have, similarly to (1.6)–(1.8),

$$\begin{aligned} \bar{\sigma}_Z P_Z &= P_Z \text{ for every } \sigma \in G, \\ \bar{\sigma}_Y P_Y(C \wedge z^{-1}(D)) &= \int_D \bar{\sigma}_Y P_Y(C:z) \bar{\sigma}_Z P_Z(dz), \quad C \in \mathfrak{C}, \quad D \in \mathfrak{D}, \end{aligned}$$

and

$$\bar{\sigma}_Y P_Y(C:z) = P_Y(\sigma_Y^{-1}C:z) \text{ for every } \sigma \in G, \\ P_Z\text{-almost everywhere on } Z.$$

Moreover,  $z_1(x) = z(y(x))$  is also a  $G$ -invariant statistic on  $X$ , and we have

$$\begin{aligned} \bar{\sigma} P(B \wedge y^{-1}(C \wedge z^{-1}(D))) \\ = \int_D P_Z(dz) \int_C \bar{\sigma} P(B:y, z) \bar{\sigma}_Y P_Y(dy:z), \end{aligned}$$

where  $\sigma P(B:y, z)$  is a conditional probability of  $B$  for given  $z_1(x) = z$  and  $y(x) = y$ .

Especially, if  $\tau \in G_0$ , we have

$$(1.9) \quad \bar{\sigma} P(\tau^{-1}B:y, z) = \bar{\sigma} P(B:y, z) \quad P_Y\text{-almost everywhere on } Y.$$

Let  $x \in X$  be fixed for a moment, and write  $I_x = \{\rho: \rho x = x\}$  and  $g_x = \{\rho x: \rho \in g\}$ . The correspondence  $G_0 \rho x \rightarrow \rho I_x G_0$  of  $\{G_0 \rho x: \rho \in G\}$  onto the left coset space  $G/I_x G_0$  is 1:1. Since  $G_0$  is a normal subgroup of each of  $G$  and  $I_x G_0$ , and since  $\bar{I}_x = \{\bar{\rho}: \bar{\rho} x = x\}$  is the factor group of  $I_x G_0$  modulo  $G_0$ , the correspondence  $\rho I_x G_0 \rightarrow \bar{\rho} \bar{I}_x$  of  $G/I_x G_0$  onto  $\bar{G}/\bar{I}_x$  is also 1:1. From this fact we may consider that  $Y_z = \{y: z(y) = z\}$  coincides with the topological space  $\bar{G}/\bar{I}_x$  for some  $x \in X_z = \{x: z_1(x) = z\}$ .

**Assumption C.** There is a  $\mathfrak{B}$ -measurable subset  $X_0$  of  $X$  such that

- 1)  $P_Z$ -almost all  $X_z$  cross with  $X_0$  at only one point  $x_z$ ,
- 2)  $J_z = \{\bar{\rho}: \bar{\rho} x_z = x_z\}$  is closed and compact in  $\bar{G}$ ,
- 3) the  $\sigma$ -field  $\mathfrak{C}_z = \{C \wedge Y_z: C \in \mathfrak{C}\}$  contains every Borel subset of  $Y_z$  which is considered as a  $\sigma$ -compact and locally compact space  $\bar{G}/\bar{J}_z$ ,
- 4) the exceptional set  $N_z \subset Y_z$  where (1.9) does not hold is independent of  $\bar{\sigma} \in \bar{G}$  and  $\tau \in G_0$ ,
- 5) there is one and only one  $G_0$ -invariant measure of total measure 1 on every  $X_y = \{x: y(x) = y\}$ .

From 4) and 5) of Assumption C,  $\bar{\sigma} P(B:y, z)$  is an invariant measure on  $X_y$ , and hence independent of  $\bar{\sigma}$ . We may write this invariant measure by  $n(B:y, z)$ . Therefore we have

$$(1.10) \quad \begin{aligned} \bar{\sigma} P(B \wedge y^{-1}(C \wedge z^{-1}(D))) \\ = \int_D P_Z(dz) \int_C n(B:y, z) \bar{\sigma}_Y P_Y(dy:z), \quad D \in \mathfrak{D}, \quad C \in \mathfrak{C}, \end{aligned}$$

and

$$(1.11) \quad \int_X f(x) \bar{\sigma}P(dx) = \int_Z P_Z(dz) \int_{Y_z} f_n(y, z) \bar{\sigma}_Y P_Y(dy; z)$$

for  $\mathfrak{B}$ -measurable function  $f(x)$ , where

$$f_n(y, z) = \int_{X_y} f(x) n(dx; y, z).$$

**Theorem 1.7.** *Under Assumptions A, B and C, every  $G_0$ -invariant statistic  $y(x)$  is sufficient for  $\Omega$ , and the mapping  $\bar{\sigma}P \rightarrow \bar{\sigma}_Y P_Y$  of  $\Omega$  onto  $\Omega^Y$  is isometric. And further  $G^Y$  is isomorphic algebraically to  $\bar{G}$  under the mapping  $\sigma_Y \rightarrow \bar{\sigma}$ .*

**Proof.** From (1.10) and (1.11) follows the sufficiency of  $y(x)$ .

Let  $\{C^+, C^-\}$  be a disjoint Hahn decomposition of  $Y$  with respect to an additive set function  $(\bar{\sigma}_Y P_Y - \bar{\tau}_Y P_Y)$ . Since

$$\begin{aligned} & \bar{\sigma}_Y P_Y(C \cap z^{-1}(D)) - \bar{\tau}_Y P_Y(C \cap z^{-1}(D)) \\ &= \int_D \{ \bar{\sigma}_Y P_Y(C; z) - \bar{\tau}_Y P_Y(C; z) \} P_Z(dz), \quad C \in \mathfrak{C}, D \in \mathfrak{D}, \end{aligned}$$

it holds for every  $C \in \mathfrak{C}$  that

$$\bar{\sigma}_Y P_Y(C^+; z) - \bar{\tau}_Y P_Y(C^+; z) \geq \bar{\sigma}_Y P_Y(C; z) - \bar{\tau}_Y P_Y(C; z) \quad P_Z\text{-almost everywhere on } Z.$$

On the other hand, writing  $B^+ = y^{-1}(C^+)$ ,

$$\begin{aligned} & \bar{\sigma}_Y P_Y(C^+; z) - \bar{\tau}_Y P_Y(C^+; z) = \bar{\sigma}P(B^+; z) - \bar{\tau}P(B^+; z) \\ & \geq \int_{C^+} n(B; y, z) \{ \bar{\sigma}_Y P_Y(dy; z) - \bar{\tau}_Y P_Y(dy; z) \} \geq \bar{\sigma}P(B; z) - \bar{\tau}P(B; z) \end{aligned}$$

for every  $B \in \mathfrak{B}$  and for  $P_Z$ -almost all  $z \in Z$ . Therefore

$$\begin{aligned} \bar{\sigma}P(B^+) - \bar{\tau}P(B^+) &= \int_Z \{ \bar{\sigma}P(B^+; z) - \bar{\tau}P(B^+; z) \} P_Z(dz) \\ &\geq \int_Z \{ \bar{\sigma}P(B; z) - \bar{\tau}P(B; z) \} P_Z(dz) = \bar{\sigma}P(B) - \bar{\tau}P(B) \end{aligned}$$

for every  $B \in \mathfrak{B}$ .

Similarly we have also, by writing  $B^- = y^{-1}(C^-)$ ,

$$\bar{\sigma}P(B^-) - \bar{\tau}P(B^-) \leq \bar{\sigma}P(B) - \bar{\tau}P(B) \text{ for every } B \in \mathfrak{B}.$$

This shows that  $\{B^+, B^-\}$  is a Hahn decomposition of  $X$  with respect to  $(\bar{\sigma}P - \bar{\tau}P)$ . Hence we have

$$V(\bar{\sigma}P, \bar{\tau}P) = V_Y(\bar{\sigma}_Y P_Y, \bar{\tau}_Y P_Y).$$

Since  $G$  is a group of transformations on  $X$ , we have

$$\bigcap_{x \in X} I_x = \{e\}, \quad I_x = \{\rho : \rho x = x\},$$

and hence

$$\bigcap_{x \in X} I_x G_0 = G_0.$$

This shows that each  $\sigma$  outside of  $G_0$  is a transformation of a coset space  $G/I_x G_0$  for some  $x$ , and that  $G^y$  is isomorphic to  $\bar{G} = G/G_0$ .

**Assumption D.** *There are a  $G$ -invariant statistic and a  $G_0$ -invariant statistic.*

By Theorem 1.7 it is sufficient for statistical problems to treat only a  $G_0$ -invariant statistic instead of the sample itself. Hence, in the remainder of this section and in Section 2, we assume without any loss of generality that  $G$  is isomorphic to  $\bar{G}$  as a topological group, and that sample space  $X$  is the union of the topological space  $X_z$  each of which is homogeneous with respect to some factor group of  $G$ . Hence we shall use, without any confusion, the same notation for elements of  $G$ ,  $\bar{G}$  and  $G^y$ .

By 2) of Assumption C, there is a unique Haar measure  $Q(L:z)$  of total measure 1 on  $J_z$ . Let  $L \in \mathfrak{L}$ . It is proved similarly as Theorem F of [10, page 281] that  $Q(\tau^{-1}L \cap J_z:z)$  is Borel measurable in  $X_z$  as a function of  $x = \tau x_z$ . Write

$$\bar{\sigma P}(L:z) = \int_{X_z} Q(\tau^{-1}L \cap J_z:z) \sigma P(d(\tau x_z):z).$$

The probability measure  $\bar{\sigma P}(L:z)$  thus defined on  $(G, \mathfrak{L})$  satisfies obviously that

$$\bar{\sigma P}(L; z) = \bar{P}(\sigma^{-1}L: z).$$

We call  $\bar{P}(L:z)$  the *fiducial measure* for  $z$ . Denoting

$$\bar{f}_z(\tau) = f(\tau x_z)$$

for a measurable function  $f(x)$ , we have

$$(1.12) \quad \int_X f(x) \sigma P(dx) = \int_Z P_z(dz) \int_G \bar{f}_z(\sigma \tau) \bar{P}(d\tau:z).$$

In fact,

$$\begin{aligned} f(\tau x_z) &= \int_{J_z} f(\tau \rho x_z) Q(d\rho:z) = \int_G f(\tau \rho x_z) Q((d\rho) \cap J_z:z) \\ &= \int_G f(\rho x_z) Q(\tau^{-1}(d\rho) \cap J_z:z) \\ &= \int_G \bar{f}_z(\rho) Q(\tau^{-1}(d\rho) \cap J_z:z), \end{aligned}$$

and

$$\begin{aligned} \int_X f(x) \sigma P(dx) &= \int_Z P_z(dz) \int_{X_z} f(x) \sigma P(dx:z) \\ &= \int_Z P_z(dz) \int_{X_z} f(\tau x_z) \sigma P(d(\tau x_z):z) \end{aligned}$$

$$= \int_Z P_Z(dz) \int_{X_z} \sigma P(d(\tau x_z):z) \int_G \bar{f}_z(\rho) Q(\tau^{-1}(d\rho) \sim J_z:z).$$

Hence by Robins' theorem [21, Th. 2]

$$\begin{aligned} &= \int_Z P_Z(dz) \int_G \bar{f}_z(\rho) \bar{\sigma} \bar{P}(d\rho:z) \\ &= \int_Z P_Z(dz) \int_G \bar{f}_z(\sigma\rho) \bar{P}(d\rho:z). \end{aligned}$$

**2. Minimax invariant decision functions.**

*2.1 Simple decision problems admitting a group of transformations.*

In Wald's theory [22] of statistical decision functions, the decision problem of fixed sample size is determined by the sample space  $X$ , the distribution space  $\Omega$ , the decision space  $\mathbf{A}$ , the space  $\Phi$  of decision functions and the loss function  $W(P', \mathbf{a}): P' \in \Omega$  and  $\mathbf{a} \in \mathbf{A}$ . In this section, we assume that

a) There is a group  $G$  of transformations on the sample space  $X$ , under each transformation  $\sigma$  of which a fixed probability measure  $P$  on  $X$  is mapped into  $\bar{\sigma}P(B) = P(\sigma^{-1}B)$ .

b) The distribution space  $\Omega$  is the set of all images  $\bar{\sigma}P$  of a fixed  $P$  under transformations  $\sigma \in G$ . Hence  $\bar{G}$  is the parameter group as defined in 1.1. The metric space  $\Omega$  with  $V$  as a metric satisfies Assumptions A and B. From Theorem 1.1  $\bar{G}$  becomes a topological group. Denote by  $\mathcal{A}$  and  $\mathcal{B}$  the classes of all Borel subsets of  $\Omega$  and  $\bar{G}$  respectively.

c)  $(\mathbf{A}, \mathcal{U})$  is a measurable space, and  $\tilde{G}$  is a group of 1:1  $\mathcal{U}$ -measurable transformations  $\tilde{\sigma}$  on  $\mathbf{A}$ , and  $\bar{G}$  is homomorphic to  $\tilde{G}$  under the mapping  $\bar{\sigma} \rightarrow \tilde{\sigma}$ .

d)  $W(\bar{\rho}\bar{\sigma}P, \bar{\rho}\mathbf{a}) = W(\bar{\sigma}P, \mathbf{a})$  holds for every  $\rho, \sigma \in G$  and for every  $\mathbf{a} \in \mathbf{A}$ .  $W(\bar{\sigma}P, \mathbf{a})$  is a non-negative and  $\mathcal{A} \times \mathcal{U}$ -measurable function on  $\Omega \times \mathbf{A}$ . Write

$$w(\mathbf{a}) = W(P, \mathbf{a}), \mathbf{a} \in \mathbf{A}.$$

e)  $\Phi$  is the class of the probability measures  $\varphi(A:x)$  on  $(\mathbf{A}, \mathcal{U})$  depending on  $x$ , such that  $\varphi(\tilde{\sigma}A:\sigma x)$  is an  $\mathcal{B} \times \mathcal{B}$ -measurable function on  $G \times X$  for any fixed  $A \in \mathcal{U}$ .  $\varphi \in \Phi$  is called a decision function.

f) Assumptions C and D are fulfilled by  $X$  and  $G$ .

A decision problem  $(X, \Omega, \mathbf{A}, W, \Phi)$  satisfying the conditions a)—f) will be called a *simple* decision problem admitting the group  $G$ .

A decision function  $\varphi(A:x)$  ( $\in \Phi$ ) satisfying the condition

$$\varphi(\tilde{\sigma}A:\sigma x) = \varphi(A:x)$$

is said to be *invariant*. The probability measure

$$D(A: \varphi, \bar{\sigma}P) = \int_x \varphi(A: x) \bar{\sigma}P(dx) = \int_x \varphi(A: \sigma x) P(dx)$$

on  $(A, \mathfrak{A})$  is called a distribution of decision under  $\varphi$  when  $\bar{\sigma}P$  is true. If  $\varphi$  is invariant, then

$$D(A: \varphi, \bar{\sigma}\bar{\rho}P) = D(\bar{\sigma}^{-1}A: \varphi, \bar{\rho}P).$$

If  $g$  is a subgroup of  $G$ , and if  $\varphi(A: x)$  is a function of  $x$  only through a  $g$ -invariant statistics for any fixed  $A$ , then  $D(A: \varphi, \bar{\sigma}P)$  may be regarded as a function of the right coset space  $\bar{G}/\bar{g}$  for any fixed  $A \in \mathfrak{A}$  and  $\varphi$ .

The risk function  $r(\bar{\sigma}P, \varphi)$  is one of the essential concepts in Wald's theory. It is defined as

$$(2.1) \quad r(\bar{\sigma}P, \varphi) = \int_x \bar{\sigma}P(dx) \int_A W(\bar{\sigma}P, \mathbf{a}) \varphi(d\mathbf{a}: x).$$

By applying the definition of  $\bar{\sigma}P$  and the condition d) to the integral (2.1) and then transforming  $\mathbf{a} \rightarrow \bar{\sigma}\mathbf{a}$  and  $x \rightarrow \sigma x$ , we have

$$r(\bar{\sigma}P, \varphi) = \int_x P(dx) \int_A w(\mathbf{a}) \varphi(\bar{\sigma}d\mathbf{a}: \sigma x).$$

From (1.12),

$$(2.2) \quad r(\bar{\sigma}P, \varphi) = \int_z P_z(dz) \int_G \bar{P}(d\tau: z) \int_A w(\mathbf{a}) \varphi_n(\bar{\sigma}d\mathbf{a}: \sigma_Y \tau_Y y_z),$$

$$y_z = y(x_z),$$

where  $z(x)$  is a  $G$ -invariant statistic, and

$$\varphi_n(A: y) = \int_{x_y} \varphi(A: x) n(dx: y, z).$$

By applying Robins' theorem [21, Th. 2] and Fubini's theorem, it follows from the condition e) that  $r(\bar{\sigma}P, \varphi)$  is a  $\mathfrak{B}$ -measurable function on  $\bar{G}$  for any fixed  $\varphi \in \Phi$ . Hence  $r(\bar{\sigma}P, \varphi)$  is also  $\mathcal{A}$ -measurable on  $\Omega$  since the mapping  $\bar{\sigma} \rightarrow \bar{\sigma}P$  is open and continuous.

For the completeness, we shall give some definitions in the theory of decision functions.

i) A probability measure  $\pi$  on  $(\Omega, \mathcal{A})$  is called an a priori distribution, and the integral of the risk function

$$r^*(\pi, \varphi) = \int_{\Omega} r(\bar{\sigma}P, \varphi) \pi(d\bar{\sigma}P)$$

is called an average risk with respect to the a priori distribution  $\pi$ .

ii) If a decision function  $\varphi_{\pi} \in \Phi$  is such that

$$r^*(\pi, \varphi_{\pi}) \leq r^*(\pi, \varphi) \text{ for any } \varphi \in \Phi,$$

then  $\varphi_{\pi}$  is called a Bayes solution relative to  $\pi$ . If, in general,  $\pi_1, \pi_2, \dots$  is a sequence of a priori distributions, and  $\varphi_0$  is a decision function  $\in \Phi$  such that

$$\lim_{n \rightarrow \infty} \{r^*(\pi_n, \varphi_0) - \inf_{\varphi \in \Phi} r^*(\pi_n, \varphi)\} = 0.$$

then  $\varphi_0$  is called a Bayes solution in the wide sense relative to the sequence  $\pi_1, \pi_2, \dots$

iii) A decision function  $\varphi_0 \in \Phi$  is said to be minimax, if

$$\sup_{\bar{\sigma}P \in \Omega} r(\bar{\sigma}P, \varphi) \geq \sup_{\bar{\sigma}P \in \Omega} r(\bar{\sigma}P, \varphi_0)$$

holds for every  $\varphi \in \Phi$ .

iv) An a priori distribution  $\pi_0$  is said to be least favourable, if

$$\inf_{\varphi \in \Phi} r^*(\pi_0, \varphi) \geq \inf_{\varphi \in \Phi} r^*(\pi, \varphi)$$

holds for every a priori distribution  $\pi$ .

v) A decision function  $\varphi \in \Phi$  is said to be non-randomized, if there exists a measurable mapping  $\mathbf{a}(x)$  of  $(X, \mathfrak{B})$  into  $(\mathbf{A}, \mathfrak{A})$  such that

$$\varphi(A; x) = \begin{cases} 1, & \text{if } A \ni \mathbf{a}(x), \\ 0, & \text{otherwise.} \end{cases}$$

vi) A decision function  $\varphi_0$  is admissible, if there is no  $\varphi \in \Phi$  such that

$$r(\bar{\sigma}P, \varphi_0) \geq r(\bar{\sigma}P, \varphi) \text{ for all } \bar{\sigma}P \in \Omega,$$

and

$$r(\bar{\sigma}P, \varphi_0) > r(\bar{\sigma}P, \varphi) \text{ for some } \bar{\sigma}P \in \Omega.$$

## 2.2. *A*-groups.

In this paragraph we shall give for the purpose of the later discussions a definition of *A*-groups, and some examples of them.

Denote by  $\mu^r$  the right invariant Haar measure of  $G$ . We shall say that a locally compact and  $\sigma$ -compact group  $G$  is an *A*-group, if corresponding to any compact set  $J \subset G$  and to a positive  $\varepsilon$  there is a compact set  $K \subset G$  such that

$$(2.3) \quad \mu^r(K) > 0 \text{ and } 1 - \frac{\mu^r(K)}{\mu^r(K \cdot J^{-1})} < \varepsilon.$$

**Theorem 2.1.** *Each of the following group is an A-group:*

- 1) *The additive group of the integrals.*
- 2) *The additive group of the real numbers.*
- 3) *A compact group.*
- 4) *The direct product group of a finite number of A-groups.*
- 5) *A locally compact, connected and commutative group.*
- 6) *The product  $G_1 \cdot G_2$  of a closed normal subgroup  $G_1$ , and a closed subgroup  $G_2$  (not necessarily normal) such that  $G_1 \cap G_2 = \{e\}$ , and that both of  $G_1$  and  $G_2$  are A-groups.*

**Proof.** 1)-4) is evident.

5) follows from Theorem 41 of [20] and 2)-4).

For 6), it is sufficient to prove that (2.3) holds for the product set  $J=J_1 \cdot J_2$  of two compact sets  $J_1 \subset G_1$  and  $J_2 \subset G_2$ . Since  $G_2$  is an  $A$ -group, we can choose a compact set  $F_2 \subset G_2$  such that

$$(1-\varepsilon)\mu_2(F_2 \cdot J_2^{-1}) < \mu_2(F_2),$$

where  $\mu_2$  is the right invariant Haar measure on  $G_2$ . Write

$$I = \bigcup_{\sigma \in F_2 \cdot J_2^{-1}} \sigma J_1 \sigma^{-1}.$$

This set  $I$ , being the image of the compact subset  $J_1 \times (F_2 \cdot J_2^{-1})$  of  $G_1 \times G_2$  by the continuous mapping  $(\sigma, \tau) \rightarrow \sigma \tau \sigma^{-1}$ , is also compact in the  $A$ -group  $G_1$ . Hence there is a compact subset  $F_1$  of  $G_1$ , such that

$$(1-\varepsilon)\mu_1(F_1 \cdot I^{-1}) < \mu_1(F_1),$$

where  $\mu_1$  is the right invariant Haar measure on  $G_1$ . By writing

$$F = F_1 \cdot F_2,$$

we have

$$\begin{aligned} \mu^r(F \cdot J^{-1}) &= \mu^r(F_1 \cdot F_2 \cdot J_2^{-1} \cdot J_1^{-1}) \\ &= \int_{F_2 \cdot J_2^{-1}} \mu_1(F_1 \cdot \sigma \cdot J_1^{-1} \cdot \sigma^{-1}) \mu_2(d\sigma) \\ &\leq \mu_1(F_1 \cdot I^{-1}) \mu_2(F_2 \cdot J_2^{-1}) \\ &< (1-\varepsilon)^{-2} \mu_1(F_1) \mu_2(F_2) \\ &= (1-\varepsilon)^{-2} \mu^r(F), \end{aligned}$$

since it holds for every  $\sigma \in F_2 \cdot J_2^{-1}$  that

$$F_1 \cdot F_2 \cdot J_2^{-1} \cdot J_1^{-1} \sigma^{-1} \cap G_1 = F_1 \cdot \sigma \cdot J_1^{-1} \cdot \sigma^{-1}.$$

In general, the question whether every  $\sigma$ -compact and locally compact group is an  $A$ -group remains open.

### 2.3 An extension of Blackwell-Girshick's theorem.

Similarly to the last part of the preceding section, we shall assume without any loss of generality that  $G = \bar{G} = G^r$ . (see page 42).

Fix a non-descending sequence  $J_1, J_2, \dots$  of compact subsets of  $G$ , tending to  $G$ , and write

$$\mathbf{b}_n(\mathbf{a}, z) = \int_{J_n} w(\tilde{\rho}\mathbf{a}) \bar{P}(d\rho: z), \quad n=1, 2, \dots,$$

and

$$\mathbf{b}(\mathbf{a}, z) = \int_G w(\rho\mathbf{a}) \bar{P}(d\rho: z),$$

where  $\bar{P}(L: z)$  is the fiducial measure for  $z$ .



**Theorem 2.2.** *Suppose that  $(X, \Omega, \mathbf{A}, W, \Phi)$  is a simple decision problem admitting an  $A$ -group  $G$ . It holds for every decision function  $\varphi \in \Phi$  that*

$$\sup_{\sigma P \in \Omega} r(\sigma P, \varphi) \geq F^*,$$

*if there is an integer  $n$  for any positive  $\varepsilon$  and for any  $z \in Z - N$  ( $P_z(N) = 0$ ) such that*

$$(2.4) \quad \mathbf{b}_n(\mathbf{a}, z) > F(z) - \varepsilon \text{ for every } \mathbf{a} \in \mathbf{A},$$

*where*

$$F(z) = \inf_{\mathbf{a} \in \mathbf{A}} \mathbf{b}(\mathbf{a}, z) \text{ and } F^* = \int_Z F(z) P_z(dz).$$

**Proof.** We assume without any loss of generality that

$$\sup_{\sigma P \in \Omega} r(\sigma P, \varphi) < \infty.$$

It follows from (2.2) and Fubini's theorem that

$$(2.5) \quad \begin{aligned} R(K, \varphi) &= \int_K r(\sigma P, \varphi) \mu^r(d\sigma) \\ &= \int_K \mu^r(d\sigma) \int_Z P_z(dz) \int_G \bar{P}(d\rho:z) \int_{\mathbf{A}} w(\mathbf{a}) \varphi(\partial d\mathbf{a}: \sigma \rho x_z) \\ &= \int_Z P_z(dz) \int_G \bar{P}(d\rho:z) \int_K \mu^r(d\sigma) \int_{\mathbf{A}} w(\mathbf{a}) \varphi(\partial d\mathbf{a}: \sigma \rho x_z) \end{aligned}$$

for every compact set  $K \subset G$ . By applying the transformation  $\mathbf{a} \rightarrow \tilde{\rho} \mathbf{a}$ , we have

$$(2.6) \quad \int_{\mathbf{A}} w(\mathbf{a}) \varphi(\partial d\mathbf{a}: \sigma \rho x_z) = \int_{\mathbf{A}} w(\tilde{\rho} \mathbf{a}) \varphi(\partial \tilde{\rho} d\mathbf{a}: \sigma \rho x_z).$$

Again by transforming  $\sigma \rightarrow \sigma \rho^{-1}$ ,

$$(2.7) \quad \int_K \varphi(\partial \tilde{\rho} A': \sigma \rho x_z) \mu^r(d\sigma) = \int_{K \cdot \rho} \varphi(\partial A': \sigma x_z) \mu^r(d\sigma)$$

for every  $A' \in \mathcal{A}$ . From (2.5), (2.6) and (2.7), and by using Robins' theorem [21, Th. 2], we have

$$(2.8) \quad R(K, \varphi) = \int_Z P_z(dz) \int_G \bar{P}(d\rho:z) \int_{K \cdot \rho} \mu^r(d\sigma) \int_{\mathbf{A}} w(\tilde{\rho} \mathbf{a}) \varphi(\partial d\mathbf{a}: \sigma x_z).$$

On the other hand, by Fubini's theorem and by the fact  $\{(\rho, \sigma): \rho \in G, \sigma \in K\rho\} = \{(\rho, \sigma): \rho \in K^{-1} \cdot \sigma, \sigma \in G\}$ ,

$$\int_G \bar{P}(d\rho:z) \int_{K \cdot \rho} f(\sigma, \rho) \mu^r(d\sigma) = \int_G \mu^r(d\sigma) \int_{K^{-1} \cdot \sigma} f(\sigma, \rho) \bar{P}(d\rho:z)$$

holds  $P_z$ -almost everywhere on  $Z$  for every  $\mathfrak{B} \times \mathfrak{B}$ -measurable function  $f(\sigma, \rho)$ . Since

$$\int_{\mathbf{A}} w(\tilde{\rho}\mathbf{a})\varphi(\tilde{\sigma}d\mathbf{a}:x_z)$$

is  $\mathfrak{Z} \times \mathfrak{Z}$ -measurable, (2.8) becomes

$$\int_{\mathfrak{Z}} P_z(dz) \int_G \mu^r(d\sigma) \int_{K^{-1} \cdot \sigma} \bar{P}(d\rho:z) \int_{\mathbf{A}} w(\tilde{\rho}\mathbf{a})\varphi(\tilde{\sigma}d\mathbf{a}:\sigma x_z).$$

Therefore from Fubini's theorem follows

$$(2.9) \quad R(K, \varphi) = \int_{\mathfrak{Z}} P_z(dz) \int_G \mu^r(d\sigma) \int_{\mathbf{A}} \varphi(\tilde{\sigma}d\mathbf{a}:\sigma x_z) \int_{K^{-1} \cdot \sigma} w(\tilde{\rho}\mathbf{a}) \bar{P}(d\rho:z).$$

Now let  $\varepsilon$  be any arbitrary positive number,  $D_n$  a  $\mathfrak{D}$ -measurable set of all  $z \in \mathfrak{Z}$  for which (2.4) holds, and  $K_n$  a compact set satisfying (2.3) for  $J=J_n$ . Writing

$$(2.10) \quad K_n^* = K_n \cdot J_n^{-1},$$

$K_n^*$  is compact and  $K_n^{*-1} \cdot \sigma \supset J_n$  holds for every  $\sigma \in K_n$ . Hence (2.9) yields

$$(2.11) \quad \begin{aligned} R(K_n^*, \varphi) &\geq \int_{D_n} P_z(dz) \int_{K_n} \mu^r(d\sigma) \int_{\mathbf{A}} \mathbf{b}_n(\mathbf{a}, z) \varphi(\tilde{\sigma}d\mathbf{a}:\sigma x_z) \\ &> \int_{D_n} P_z(dz) \int_{K_n} \mu^r(d\sigma) \int_{\mathbf{A}} [F(z) - \varepsilon] \varphi(\tilde{\sigma}d\mathbf{a}:\sigma y_z) \\ &= \mu^r(K_n) \int_{D_n} [F(z) - \varepsilon] P_z(dz) \\ &> (1 - \varepsilon) \mu^r(K_n^*) \int_{D_n} [F(z) - \varepsilon] P_z(dz). \end{aligned}$$

Evidently  $\{D_n\}$  is a non-descending sequence and tends to  $\mathfrak{Z} - N$ , since  $\{J_n\}$  is non-descending. Hence

$$\lim_{n \rightarrow \infty} \int_{D_n} F(z) P_z(dz) = \int_{\mathfrak{Z}} F(z) P_z(dz) = F^*.$$

That is to say, there is an integer  $n_0$  such that

$$(2.12) \quad \int_{D_n} F(z) P_z(dz) < F^* - \varepsilon \text{ for every } n > n_0.$$

Hence we have, from (2.11) and (2.12),

$$(2.13) \quad \sup_{\sigma P \ni \Omega} r(\sigma P, \varphi) \geq \frac{R(K_n^*, \varphi)}{\mu^r(K_n^*)} > (1 - \varepsilon)(F^* - 2\varepsilon),$$

where  $n$  is a larger integer than  $n_0$ .

Since  $\varepsilon$  is arbitrary, (2.13) shows that

$$\sup_{\sigma P \in \Omega} r(\sigma P, \varphi) \geq F^*.$$

Thus the proof is complete.

This theorem is a straightforward extension of Blackwell-Girshick's result, and the proof developed above has been proceeded exactly along the way of theirs.

**Corollary.** Under the hypotheses of Theorem 2.2, the invariant decision function  $\varphi^0$  is minimax, if and only if

$$(2.14) \quad \varphi^0(A_z : x_z) = 1 \quad P_z\text{-almost everywhere on } Z,$$

provided that  $A_z = \{\mathbf{a} : \mathbf{b}(\mathbf{a}, z) = F(z)\}$  is not empty.

**Proof.** For  $\varphi^0$  satisfying (2.14) it follows from (2.2) and Fubini's theorem that

$$(2.15) \quad \begin{aligned} r(\sigma P, \varphi^0) &= \int_Z P_z(dz) \int_G \bar{P}(d\tau : z) \int_A w(\mathbf{a}) \varphi^0(d\mathbf{a} : \sigma\tau x_z) \\ &= \int_Z P_z(dz) \int_G \bar{P}(d\tau : z) \int_A w(\tilde{\tau}\mathbf{a}) \varphi^0(d\mathbf{a} : x_z) \\ &= \int_Z P_z(dz) \int_A \mathbf{b}(\mathbf{a}, z) \varphi^0(d\mathbf{a} : x_z) \\ &= \int_Z P_z(dz) \int_{A_z} \mathbf{b}(\mathbf{a}, z) \varphi^0(d\mathbf{a} : x_z) \\ &= \int_Z P_z(dz) \int_{A_z} F(z) \varphi^0(d\mathbf{a} : x_z) \\ &= \int_Z F(z) P_z(dz) = F^*. \end{aligned}$$

Hence  $\varphi^0$  is minimax by Theorem 2.2.

Denoting by  $K_n$  a compact subset of  $G$ , satisfying (2.3) for  $J = J_n$ , we have, similarly to (2.11),

$$\begin{aligned} \frac{R(K_n \cdot J_n^{-1}, \varphi)}{\mu^r(K_n \cdot J_n^{-1})} &\geq \frac{\mu^r(K_n)}{\mu^r(K_n \cdot J_n^{-1})} \int_Z P_z(dz) \int_A \mathbf{b}_n(\mathbf{a}, z) \varphi(d\mathbf{a} : x_z) \\ &> (1 - \varepsilon) \left[ F^* + \int_Z P_z(dz) \int_A \{\mathbf{b}_n(\mathbf{a}, z) - F(z)\} \varphi(d\mathbf{a} : x_z) \right] \end{aligned}$$

for every invariant decision function  $\varphi$ . If  $\varphi$  is minimax, then

$$F^* = \sup_{\sigma P \in \Omega} r(\sigma P, \varphi) \geq \frac{R(K_n \cdot J_n^{-1}, \varphi)}{\mu^r(K_n \cdot J_n^{-1})}.$$

Hence for a minimax and invariant  $\varphi$  we have

$$\frac{\varepsilon}{1 - \varepsilon} F^* > \int_Z P_z(dz) \int_A \{\mathbf{b}_n(\mathbf{a}, z) - F(z)\} \varphi(d\mathbf{a} : x_z),$$

( $n = 1, 2, \dots$  ad. inf.)

Since  $\lim \mathbf{b}_n(\mathbf{a}, z) = \mathbf{b}(\mathbf{a}, z) \geq F(z)$ , it follows that

$$\int_Z P_z(dz) \int_A \{\mathbf{b}(\mathbf{a}, z) - F(z)\} \varphi(d\mathbf{a}; x_z) = 0,$$

and

$$\int_A \{\mathbf{b}(\mathbf{a}, z) - F(z)\} \varphi(d\mathbf{a}; x_z) = 0 \quad P_z\text{-almost everywhere on } Z.$$

Hence we have

$$\varphi(A_z, x_z) = 1 \quad P_z\text{-almost everywhere on } Z.$$

This Collolary shows that the minimax invariant decision function can be constructed if and only if  $\mathbf{b}(\mathbf{a}, z)$  attains its minimal value for  $P_z$ -almost every  $z \in Z$ . In general, a minimax decision function is not necessarily invariant. We shall speak of a minimax invariant decision function defined above as a *Pitman-Girshick-Savage-Blackwell's decision function* (for shorten, *PGSB d.f.*).

The following theorem is also similar to Blackwell-Girshick's remark [4, p. 311].

**Theorem 2.3.** *The hypothesis (2.4) of Theorem 2.2 holds in each of the following cases:*

- (1)  $w(\mathbf{a})$  is bounded.
- (2) The decision space  $\mathbf{A}$  is a topological space which has a sequence  $\{C_n\}$  of compact subsets such that

$$\bigcup_{n=1}^{\infty} C_n = \mathbf{A} \quad \text{and} \quad \liminf_{n \rightarrow \infty} \inf_{\mathbf{a} \notin C_n} w(\mathbf{a}) = \infty,$$

and moreover  $w(\mathbf{a})$  is a continuous function of  $\mathbf{a} \in \mathbf{A}$ .

The proof is quite similar to that of Blackwell-Girshick.

#### 2.4. Miscellaneous properties of PGSB d.f.

**Theorem 2.4.** *If  $F^* < \infty$ , then a PGSB d.f.  $\varphi^0$  is a Bayes solution in the wide sense relative to a sequence  $\{\xi_n\}$ :*

$$\xi_n(E) = \frac{\mu^r(K_n^* \cap L_E)}{\mu^r(K_n^*)} \quad \text{for every } \Lambda\text{-measurable } E \subset \Omega, \text{ where } L_E =$$

$\{\sigma : \sigma P \in E\}$  and  $K_n^*$  is a compact subset of  $G$  defined by (2.3), (2.4) and (2.10) for  $J = J_n$  and  $\varepsilon = \varepsilon_n \downarrow 0$ .

**Proof.** From (2.13) and (2.15) follows that,

$$\begin{aligned} & \int_{\Omega} r(\sigma P, \varphi) \xi_n(d\sigma P) - \int_{\Omega} r(\sigma P, \varphi^0) \xi_n(d\sigma P) \\ &= \frac{1}{\mu^r(K_n^*)} \left[ \int_{K_n^*} r(\sigma P, \varphi) \mu^r(d\sigma) - \int_{K_n^*} r(\sigma P, \varphi^0) \mu^r(d\sigma) \right] \\ &> (1 - \varepsilon_n)(F^* - 2\varepsilon_n) - F^* = -\varepsilon_n(F^* + 2 - 2\varepsilon_n) \end{aligned}$$

hold for every  $\varphi \in \Phi$ . Hence we have

$$\lim_{n \rightarrow \infty} \left\{ \inf_{\varphi \in \Phi} \int_{\Omega} r(\sigma P, \varphi) \xi_n(d\sigma P) - \int_{\Omega} r(\sigma P, \varphi^0) \xi_n(d\sigma P) \right\} = 0.$$

**Corollary.** If  $\Omega$  is compact, a PGSB d.f.  $\varphi^0$  is a Bayes solution relative to a normalized invariant distribution  $\xi$  on  $\Omega$ , which is least favourable.

**Proof.**  $G$  is compact, since  $\Omega$  is so. Hence, letting  $K_n = G$ , (2.13) yields

$$\int_{\Omega} r(\sigma P, \varphi) \xi(d\sigma P) \geq F^* = \int_{\Omega} r(\sigma P, \varphi^0) \xi(d\sigma P)$$

for every  $\varphi \in \Phi$ , because  $G$  is unimodular and

$$\mu^r(\{\sigma : \sigma P \in E\}) = \xi(E).$$

Moreover we have

$$\begin{aligned} \inf_{\varphi \in \Phi} \int_{\Omega} r(\sigma P, \varphi) \pi(d\sigma P) &\leq \int_{\Omega} r(\sigma P, \varphi^0) \pi(d\sigma P) \\ &= F^* = \inf_{\varphi \in \Phi} \int_{\Omega} r(\sigma P, \varphi) \xi(d\sigma P). \end{aligned}$$

for any a priori distribution  $\pi$ , since  $r(\sigma P, \varphi^0)$  has a constant value  $F^*$  for every  $\sigma P \in \Omega$ .

The admissibility of the PGSB d.f.  $\varphi^0$  remains unsolved in the general case. Blackwell [2] has proved this in a special case. We have the other result, stated below, than Blackwell's.

**Theorem 2.5.** *If  $G$  is compact and if  $W(\sigma P, \mathbf{a})$  is continuous on  $\Omega$  uniformly in  $\mathbf{a} \in \mathbf{A}$  and bounded on  $\Omega \times \mathbf{A}$ , then the PGSB d.f. is admissible.*

**Proof.** Since

$$\begin{aligned} &|r(\sigma P, \varphi) - r(\tau P, \varphi)| \\ &\leq \left| \int_{\Omega} \sigma P(dx) \int_{\mathbf{A}} W(\sigma P, \mathbf{a}) \varphi(d\mathbf{a}:x) - \int_{\Omega} \tau P(dx) \int_{\mathbf{A}} W(\sigma P, \mathbf{a}) \varphi(d\mathbf{a}:x) \right| \\ &\quad + \int_{\mathbf{A}} \tau P(dx) \int_{\mathbf{A}} |W(\sigma P, \mathbf{a}) - W(\tau P, \mathbf{a})| \varphi(d\mathbf{a}:x), \end{aligned}$$

we have the continuity of the function  $r(\sigma P, \varphi)$  on  $\Omega$ . Hence if  $r(\sigma P, \varphi^0) \geq r(\sigma P, \varphi)$  for all  $\sigma P$  and if  $r(\sigma P, \varphi^0) > r(\sigma P, \varphi)$  for some  $\sigma P$ , then

$$F^* = r^*(\xi, \varphi^0) > r^*(\xi, \varphi),$$

because every open subset of  $\Omega$  is of  $\xi$ -measure positive. This contradicts with the fact that  $\varphi^0$  is a Bayes solution relative to  $\xi$ . (cf. Blyth [6, p. 28])

### 2.5. Some related problems.

a) *Coset estimates.* Let  $g$  be a closed subgroup of  $G$ . Suppose that  $\mathbf{A}$  is a right coset space modulo  $g$ , and that the group  $\tilde{G}$  of transformations of  $\mathbf{A}$  consists of the transformations  $\tilde{\rho}: g\sigma \rightarrow g\sigma\rho$ . We have then

$$\mathbf{b}(g\sigma, z) = \int_G w(g\sigma\rho) \bar{P}(d\rho: z).$$

If  $w(g\sigma)$  satisfies the condition (1) or (2) of Theorem 2.3, and if  $\sigma_z$  minimizes  $\mathbf{b}(g\sigma, z)$ , then the decision function  $\varphi^0$ :

$$(2.16) \quad \varphi^0(g\sigma_z\rho: \rho x_z) = \varphi^0(g\sigma_z: x_z) = 1$$

is a PGSB d.f. Under this decision function  $\varphi^0$ , the distribution of decision when  $P$  is true is

$$\begin{aligned} D(A: \varphi^0, P) &= \int_Z P_Z(dz) \int_G \varphi^0(A: \rho x_z) \bar{P}(d\rho: z) \\ &= \int_Z P_Z(dz) \int_G \varphi^0(\bar{\rho}^{-1}A: x_z) \bar{P}(d\rho: z) \\ &= \int_Z \bar{P}(\{\rho: A \ni g\sigma_z\rho\}: z) P_Z(dz), \end{aligned}$$

where  $A$  is a Borel set of the right coset space  $G/g$ . Hence we have

**Theorem 2.6.** *If the decision space  $\mathbf{A}$  is a right coset space  $G/g$ , then the distribution of decision under a non-randomized PGSB d.f. given as (2.16) when  $\sigma P$  is true is independent of the parameters belonging to  $\bigcap_{z \in Z} \sigma_z^{-1}g\sigma_z$  i.e.,*

$$D(A: \varphi^0, \sigma\tau P) = D(A: \varphi^0, \tau P)$$

for every  $\sigma \in \bigcap_{z \in Z} \sigma_z^{-1}g\sigma_z$ .

b) *The loss function for problems of estimation by intervals.*

Suppose that the fiducial measure  $\bar{P}$  for each  $z$  is absolutely continuous with respect to the left invariant Haar measure  $\mu$  on  $G$ , and write

$$\bar{P}(L: z) = \int_L \bar{p}(\sigma: z) \mu(d\sigma).$$

Let the decision space  $\mathbf{A}$  be the class of the subsets  $\mathbf{a}$  of  $G$  transformed from the sets of the form  $\{\sigma^{-1}: p(\sigma; z) \geq u\}$  by elements  $\tau \in G$ . Put, for  $\mathbf{a} \in \mathbf{A}$ ,

$$C(\mathbf{a}) = \begin{cases} 0 & \text{if } e \in \mathbf{a}, \\ 1 & \text{if } e \notin \mathbf{a}. \end{cases}$$

**Theorem 2.7.** *If*

$$w(\mathbf{a}) = \alpha C(\mathbf{a}) + \beta \mu(\mathbf{a}^{-1}), \quad \alpha, \beta > 0,$$

*then the invariant decision function  $\varphi$ :*

$$\varphi(\mathbf{a}_z: x_z) = 1, \quad \mathbf{a}_z = \left\{ \sigma^{-1}: \bar{p}(\sigma: z) \geq \frac{\beta}{\alpha} \right\},$$

*is a PGSB d.f.*

**Proof.**  $\mathbf{b}(\mathbf{a}, z) = \alpha \bar{P}(G - \mathbf{a}^{-1}: z) + \beta \mu(\mathbf{a}^{-1}) = \alpha + \int_{\mathbf{a}^{-1}} \{\beta - \alpha \bar{p}(\sigma: z)\} \mu(d\sigma),$

$\mathbf{b}_n(\mathbf{a}, z) = \alpha \bar{P}(J_n - \mathbf{a}^{-1}: z) + \beta \mu(\mathbf{a}^{-1}) \bar{P}(J_n: z) = \alpha \bar{P}(J_n: z) + \beta \mu(\mathbf{a}^{-1} - J_n) \bar{P}(J_n: z) + \int_{\mathbf{a}^{-1} \cap J_n} \{\beta \bar{P}(J_n: z) - \alpha \bar{p}(\sigma: z)\} \mu(d\sigma)$ . Denote  $\mathbf{a}_{n,z} = \left\{ \sigma^{-1}: \bar{p}(\sigma: z) \geq \frac{\beta}{\alpha} \bar{P}(J_n: z) \right\} \cap J_n^{-1}$ . Then obviously  $\mathbf{a}_z$  and  $\mathbf{a}_{n,z}$  minimize  $\mathbf{b}(\mathbf{a}, z)$  and  $\mathbf{b}_n(\mathbf{a}, z)$  respectively. From  $\infty > \mu \left[ \left\{ \sigma: \bar{p}(\sigma: z) > \frac{\beta}{\alpha} \cdot \bar{P}(J_n: z) \right\} \right] \geq \mu(\mathbf{a}_z^{-1})$ ,  $\lim \bar{P}(J_n: z) = 1$  and  $\bigcup_{n=1}^{\infty} \mathbf{a}_{n,z} = \mathbf{a}_z$ , it follows that there is for every  $\varepsilon > 0$  an integer  $n_0$  such that  $\mu(\mathbf{a}_{n,z}^{-1}) \bar{P}(J_n: z) > \mu(\mathbf{a}_z^{-1}) - \varepsilon$  and  $\bar{P}(J_n - \mathbf{a}_{n,z}^{-1}: z) > \bar{P}(G - \mathbf{a}_z^{-1}: z) - \varepsilon$  hold for  $n > n_0$ . This shows that (2.4) holds and by Theorem 2.2 the PGSB d.f. exists. Hence by Corollary of Theorem 2.2 the proof is accomplished.

### 3. The relation between topologies of the sample space and of the distribution space.

#### 3.1. Homogeneous spaces with respect to a group.

Suppose that i) the sample space  $X$  is a  $\sigma$ -compact, connected and locally compact metric space, where the metric in  $X$  is denoted by  $((x, y))$ ;

- ii)  $\mathfrak{B}$  is the class of all Borel sets of  $X$ ;
- iii)  $G$  is a group of isometric transformations of  $X$  onto itself;
- iv)  $X$  is homogeneous with respect to  $G$ ;
- v) if  $\sigma_1 x, \sigma_2 x, \dots$  is a Cauchy sequence of sample points for every  $x \in X$ , then there exists an element  $\sigma_0 \in G$  such that

$$\lim_{n \rightarrow \infty} \sigma_n x = \sigma_0 x \text{ for every } x \in X.$$

As seen in Theorems 1.1 and 1.3,  $G$  becomes a  $\sigma$ -compact and locally compact topological group with the base of neighbourhoods  $u(k, \varepsilon) = \{\sigma: ((\sigma x, x)) < \varepsilon \text{ for every } x \in k\}$  of the neutral element  $e$ , where  $k$  is a compact subset of  $X$ , and there is one and only one  $G$ -invariant measure  $m$  on  $(X, \mathfrak{B})$ .

Further we suppose that

- vi) the probability measure  $P(B)$  is absolutely continuous with respect to the  $G$ -invariant measure  $m$  stated above, and

$$P(B) = \int_B p(x) m(dx), B \in \mathfrak{B}.$$

From vi) every  $\bar{\sigma}P$  ( $\sigma \in G$ ) is also absolutely continuous with respect to  $m$  and we have

$$\bar{\sigma}P(B) = \int_B p(\sigma^{-1}x) m(dx), B \in \mathfrak{B}.$$

**Lemma 3.1.**  $V(\bar{\sigma}\bar{\tau}P, \bar{\tau}P)$  is a continuous function of  $\sigma$  for every fixed  $\bar{\tau} \in \bar{G}$ .

**Proof.** Choose a compact subset  $k$  of  $X$  such that

$$\bar{\tau}P(X-k) < \varepsilon/3$$

for a given positive number  $\varepsilon$ . Such a set exists, for  $X$  is  $\sigma$ -compact and  $\bar{\tau}P$  is finite measure. Let  $v(x_0, 2\delta)$  be a neighbourhood of  $x_0$  with radius  $2\delta$ , whose closure is compact. Since every transformation in  $G$  is isometric, the closure of every open set  $v(x, 2\delta)$  is compact (from the condition iv)). Denote by  $x_1, x_2, \dots, x_n$  a finite  $\delta$ -net covering  $k$ , and write  $k' = \{x_1, x_2, \dots, x_n\}$  and

$$k'' = \text{the closure of } \bigcup_{j=1}^n v(x_j, 2\delta).$$

Then we have

$$u(k', \delta)k \subset k'',$$

because there is, for any  $x \in k$  and any  $\rho \in u(k', \delta)$ , an  $x_j \in k'$  such that

$$\begin{aligned} ((\rho x, x_j)) &\leq ((\rho x, \rho x_j)) + ((\rho x_j, x_j)) \\ &= ((x, x_j)) + ((\rho x_j, x_j)) < \delta + \delta = 2\delta. \end{aligned}$$

From this fact, it holds for every  $\rho \in u(k', \delta)$  that

$$\begin{aligned} (3.1) \quad \bar{\rho}\bar{\tau}P(X-k'') &= \bar{\tau}P(X-\rho^{-1}k'') \\ &\leq \bar{\tau}P(X-k) < \varepsilon/3. \end{aligned}$$

As seen in ix) of Theorem 1.4, we can choose a compact subset  $k^*$  of  $X$  and a positive number  $\delta_1$  such that

$$(3.2) \quad [\text{absolute variation of } (\bar{\sigma}\bar{\tau}P - \bar{\tau}P) \text{ on } k''] < \varepsilon/3.$$

for every  $\sigma \in u(k^*, \delta_1)$ . Therefore it follows from (3.1) and (3.2) that

$$\begin{aligned} &V(\bar{\sigma}\bar{\tau}P, \bar{\tau}P) \\ &\leq [\text{absolute variation of } (\bar{\sigma}\bar{\tau} - \bar{\tau}P) \text{ on } k''] + \bar{\sigma}\bar{\tau}P(X-k'') + \bar{\tau}P(X-k'') \\ &< \frac{\varepsilon}{3} + \frac{2}{3}\varepsilon = \varepsilon. \end{aligned}$$

for every  $\sigma \in u(k^*, \delta_1) \cap u(k', \delta)$ . This completes the proof.

From Lemma 3.1. directly follows

**Lemma 3.2.** *The mapping  $\sigma \rightarrow \bar{\sigma}P$  of  $G$  onto  $\Omega$  is continuous. Hence  $\Omega$  is a  $\sigma$ -compact metric space.*

Hence we have, by using the remark in the footnote 8),

**Lemma 3.3.** *There corresponds for every compact subset  $K$  of  $\Omega$  and every real  $\varepsilon > 0$  a neighbourhood  $u$  of  $e$  of  $G$  such that*

$$\begin{aligned} \bar{u} &= \{\bar{\sigma} : \sigma \in u\} \\ &\subset U(K; \varepsilon) = \{\bar{\sigma} : V(\bar{\sigma}\bar{\rho}P, \bar{\rho}P) < \varepsilon \text{ for every } \bar{\rho}P \in K\}. \end{aligned}$$

**Remark 3.1.** If we have known that  $\Omega$  is locally compact, then the class of all  $U(K, \varepsilon)$ 's forms a complete system of neighbourhoods of  $\bar{e}$ , as seen in Theorem 1.1. In this case Lemma 3.3 asserts the continuity of the algebraic homomorphism  $\sigma \rightarrow \bar{\sigma}$ , and hence, moreover, it asserts



that  $\sigma \rightarrow \bar{\sigma}$  is topologically a homomorphism. But we never know, up to this time, whether  $\Omega$  is locally compact. Fortunately, however, we can prove the local compactness of  $\Omega$ . We shall show this fact.

**Lemma 3.4.** *If  $G_0 = \{e\}$ , then the class of all  $U(K, \varepsilon)$ 's forms a complete system of neighbourhoods of  $e$  of  $G$ , which defines an equivalent topology to the original one of  $G$ .*

**Proof.** It is sufficient to show that the image  $\bar{u}$  of every neighbourhood  $u$  of  $e$  contains at least one  $U(K, \delta)$ . We shall here show that there are a compact subset  $K$  and  $\delta > 0$  for every  $\varepsilon > 0$  and every  $x_0 \in X$  such that

$$U(K, \delta) \subset \{\bar{\sigma} : ((\sigma x_0, x_0)) < \varepsilon\}.$$

This is sufficient for the proof (see the footnote 8)). Let us now suppose that there exists a sequence  $\sigma_1, \sigma_2, \dots$  of elements of  $G$  such that

$$((\sigma_n x_0, x_0)) > \varepsilon \text{ and } V(\overline{\sigma_n \rho P}, \bar{\rho P}) < \frac{1}{n} \text{ for } \bar{\rho P} \in K_n \text{ and } n = 1, 2, \dots, \text{ where}$$

$K_1 \subset K_2 \subset \dots$ , and  $\lim K_n = \Omega$ . Any subsequence of  $\sigma_1, \sigma_2, \dots$  has no limiting point in  $G$ . In fact, if a subsequence  $\sigma_{n_1}, \sigma_{n_2}, \dots$  tends to  $\sigma_0$  in  $G$ , there is an integer  $j$  for every  $n$  such that  $\overline{\sigma_0^{-1} \sigma_{n_i}} \in U\left(K_n, \frac{1}{n}\right)$  for every

$i > j$  (Lemma 3.2). Hence we have  $V(\overline{\sigma_0 \rho P}, \bar{\rho P}) \leq V(\overline{\sigma_0 \rho P}, \overline{\sigma_{n_i} \rho P}) + V(\overline{\sigma_{n_i} \rho P}, \bar{\rho P}) < \frac{2}{n}$  for  $\bar{\rho P} \in K_n$ . From  $\bigcup K_n = \Omega$  follows that  $V(\overline{\sigma_0 \rho P}, \bar{\rho P}) = 0$  for every  $\bar{\rho P} \in \Omega$ . This means  $\bar{\sigma}_0 = \bar{e}$ , and hence  $\sigma_0 = e$ , since  $G_0 = \{e\}$ . This contradicts with  $((\sigma_n x_0, x_0)) > \varepsilon$ .

Let  $v$  be an open set  $\{y : ((y, x_0)) < \varepsilon'\}$ , whose closure is compact, and write  $v_1 = \{y : ((y, x_0)) < \varepsilon'/3\}$ . Since  $X$  is separable, there is an element  $\bar{\sigma}_0 P$  of  $\Omega$  for which

$$a = \bar{\sigma}_0 P(v_1) > 0.$$

Let  $n_1, n_2, \dots$  be a monotone increasing sequence of integers such that, for  $j = 1, 2, \dots$ ,

$$\frac{1}{n_j} < \frac{j}{j+1} a, \quad \bar{\sigma}_0 P \in K_{n_j},$$

and

$$\sigma_{n_j}^{-1} \in \bigcup_{i=0}^{j-1} u(\{\sigma_{n_i}^{-1} x_0\}, \varepsilon') \sigma_{n_i}^{-1},$$

where  $\sigma_{n_0} = e$ . Such a subsequence exists, since any subsequence of  $\{\sigma_n\}$  has no limiting point and since the closure of  $\sigma v = u(\{\sigma x_0\}, \varepsilon') \sigma x_0$  is compact for every  $\sigma \in G$ . For this subsequence  $\{\sigma_{n_j}\}$  hold

$$\begin{aligned} \bar{\sigma}_0 P(\sigma_{n_j}^{-1} v_1) &= \overline{\sigma_{n_j} \sigma_0 P}(v_1) > \bar{\sigma}_0 P(v_1) - V(\overline{\sigma_{n_j} \sigma_0 P}, \bar{\sigma}_0 P) \\ &> \frac{a}{j+1}, \end{aligned}$$

and

$$\sigma_{n_j}^{-1}v_1 \cap \left[ \bigcup_{i=0}^{j-1} \sigma_{n_i}^{-1}v_1 \right] = \text{empty}.$$

This gives a contradiction:

$$\begin{aligned} 1 = \bar{\sigma}_0 P(X) &\geq \bar{\sigma}_0 P\left(\bigcup_{j=0}^{\infty} \sigma_{n_j}^{-1}v_1\right) = \sum_{j=0}^{\infty} \bar{\sigma}_0 P(\sigma_{n_j}^{-1}v_1) \\ &> a \sum_{j=0}^{\infty} \frac{1}{j+1} = \infty. \end{aligned}$$

From Lemma 3.3 and 3.4 follows that, if  $G_0 = \{e\}$ , then  $\bar{G}$  is a topological group and the algebraical isomorphism  $\sigma \rightarrow \bar{\sigma}$  of  $G$  onto  $\bar{G}$  is a homeomorphism.

**Lemma 3.5.** *The subgroup  $H = \{\sigma : \bar{\sigma}P = P\}$  and hence  $G_0$ , being the meet of all  $\rho H \rho^{-1}$ , are closed in  $G$ .*

**Proof.** Let  $\tau$  be an element of the closure of  $H$ . Then, by Lemma 3.3, there is a neighbourhood  $u$  of  $e$  for every given real  $\varepsilon > 0$  such that  $U(\{\bar{\tau}P\}, \varepsilon) \supset \bar{u}$ . Since  $\tau u$  intersects with  $H$ ,  $\bar{\tau}u$  and hence  $\bar{\tau}U(\{\bar{\tau}P\}, \varepsilon)$  must intersect with  $\bar{H}$ . Let  $\bar{\sigma}$  belong both of  $\bar{H}$  and  $\bar{\tau}U(\{\bar{\tau}P\}, \varepsilon)$ . For such  $\bar{\sigma}$ , we have  $\varepsilon > V(\bar{\sigma}P, \bar{\tau}P) = V(P, \bar{\tau}P)$ . Since  $\varepsilon$  is arbitrary,  $\bar{\tau}P = P$ , i.e.  $\bar{\tau} \in H$ . The closure of  $G_0$  is evident.

By Lemma 3.5,  $\Omega$ , being a coset space of  $\bar{G}$ , is  $\sigma$ -compact and locally compact, that is to say,  $\Omega$  satisfies Assumption A, when  $G_0 = \{e\}$ .

**Lemma 3.6.**  *$\Omega$  and  $\bar{G}$  satisfy Assumption B, if  $G_0 = \{e\}$ .*

**Proof.** Suppose that  $\bar{\sigma}_1 \rho P, \bar{\sigma}_2 \rho P, \dots$  is a Cauchy sequence in  $\Omega$  for every fixed  $\bar{\rho}P$ . Since  $\Omega$  is complete, there is a limiting point  $\varphi(\bar{\rho}P)$  of  $\bar{\sigma}_1 \rho P, \bar{\sigma}_2 \rho P, \dots$  for every  $\bar{\rho}P$ .  $\varphi$  is evidently an isometric transformation of  $\Omega$  onto itself. Denote by  $K$  the group of all isometric transformations of  $\Omega$ . Since  $(\Omega, K)$  satisfies Assumptions A and B,  $K$  becomes a complete topological group by Mibu's Theorems 1.1 and 1.2, and  $\bar{G}$  is a closed subgroup of  $K$ , since  $G$  is complete. Hence there is an element  $\bar{\sigma}_0$  of  $\bar{G}$  such that  $\varphi = \bar{\sigma}_0$ . This shows that  $\lim \bar{\sigma}_n \rho P = \bar{\sigma}_0 \rho P$  for every  $\bar{\rho} \in \bar{G}$ .

The subgroup  $I$  of the elements which remains a fixed sample point  $x_0$  invariant is closed in  $G$ , and  $X$  is homeomorphic to the left coset space  $G/I$  under the mapping  $\sigma x_0 \rightarrow \sigma I$ , which can be proved similarly to Theorem 1.3. From this fact follows that the mapping  $\sigma \rightarrow \sigma x_0$  of  $G$  onto  $X$  is open, and that  $X_e = G_0 x_0$  is closed in  $X$ . Hence  $X_e$  is complete since  $X$  is complete.

**Lemma 3.7.**  *$X_e = G_0 x_0$  is compact.*

**Proof.** Let  $v(x_0, \varepsilon) = \{y : ((y, x_0)) < \varepsilon\}$  be a neighbourhood of  $x_0$ , whose closure is compact. If there exists an infinite sequence  $x_0, x_1, x_2, \dots$  in  $X_e$ , for which  $((x_i, x_j)) > \varepsilon (i \neq j; i, j = 0, 1, 2, \dots)$ , then we have a disjoint

sequence of open sets  $v(x_i, \varepsilon/2)$ ,  $i=0, 1, 2, \dots$ , each term of which satisfies

$$\begin{aligned} \bar{\sigma}P(v(x_i, \varepsilon/2)) &= \bar{\sigma}P(\tau_i v(x_0, \varepsilon/2)) \\ &= \bar{\sigma}P(v(x_0, \varepsilon/2)) \end{aligned}$$

where  $\bar{\sigma} \in \bar{G}$ ,  $x_i = \tau_i x_0$  and  $\tau_i \in G_0$ . On the other hand, there is a  $\sigma_0 \in G$  such that

$$\sigma_0 P(v(x_0, \varepsilon/2)) = P(\sigma_0^{-1} v(x_0, \varepsilon/2)) = a > 0,$$

since  $X$  is separable. This implies the contradiction:

$$1 = \bar{\sigma}_0 P(X) \geq \bar{\sigma}_0 P\left(\bigcup_{i=1}^{\infty} v(x_i, \varepsilon/2)\right) = \sum_{i=1}^{\infty} a = \infty.$$

Thus we see that  $X_\varepsilon$  has no infinite  $\varepsilon$ -net for any small  $\varepsilon > 0$ , that is to say,  $X_\varepsilon$  is totally bounded. As seen above,  $X_\varepsilon$  is complete, and hence compact.

**Lemma 3.8.** *Each  $X_\sigma = \sigma G_0 x_0$  is homeomorphic to  $X_\varepsilon$  under the mapping  $\tau x_0 \rightarrow \sigma \tau x_0$  ( $\tau \in G_0$ ).*

**Lemma 3.9.** *The subgroup  $G_0 I$  is closed in  $G$ .*

**Proof.** Let  $\sigma_0 \notin G_0 I$ . Then  $\sigma_0 x_0 \notin X_\varepsilon$ . Hence there is a neighbourhood  $v(\sigma_0 x_0, \varepsilon) = \{y : ((y, \sigma_0 x_0)) < \varepsilon\}$  disjoint with  $X_\varepsilon$ . This shows that a neighbourhood  $u(\{\sigma_0 x_0\}, \varepsilon)_{\sigma_0}$  of  $\sigma_0$  does not intersect with  $G_0 I$ .

Write

$$\begin{aligned} ((X_\sigma, X_{\sigma'})) &= \inf_{x \in X_\sigma, x' \in X_{\sigma'}} ((x, x')) \\ &= \inf_{\tau, \tau' \in G_0} ((\sigma \tau x_0, \sigma' \tau' x_0)) \\ &= \inf_{\tau \in G_0} ((x_0, \sigma^{-1} \sigma' \tau x_0)). \end{aligned}$$

Obviously we have i)  $((X_\sigma, X_{\sigma'})) \geq 0$  with equality if and only if  $\sigma^{-1} \sigma' \in G_0 I$ , ii)  $((X_\sigma, X_{\sigma'})) = ((X_{\sigma'}, X_\sigma))$ , and iii)  $((X_\sigma, X_{\sigma'}) + ((X_\sigma, X_{\sigma''})) \geq ((X_{\sigma'}, X_{\sigma''}))$ . We shall prove only iii). Given any real  $\varepsilon > 0$ , there correspond two elements  $\tau$  and  $\tau' \in G_0$  such that

$$((X_{\sigma'}, X_\sigma)) + \frac{\varepsilon}{2} > ((x_0, \sigma^{-1} \sigma' \tau x_0))$$

and

$$((X_{\sigma''}, X_\sigma)) + \frac{\varepsilon}{2} > ((x_0, \sigma^{-1} \sigma'' \tau' x_0)).$$

Hence  $((X_{\sigma'}, X_\sigma)) + ((X_\sigma, X_{\sigma''})) + \varepsilon > ((x_0, \sigma^{-1} \sigma' \tau x_0)) + ((x_0, \sigma^{-1} \sigma'' \tau' x_0)) \geq ((\sigma^{-1} \sigma' \tau x_0, \sigma^{-1} \sigma'' \tau' x_0)) = ((\sigma' \tau x_0, \sigma'' \tau' x_0)) \geq ((X_{\sigma'}, X_{\sigma''}))$ .

This metric  $((X_\sigma, X_{\sigma'}))$  defines a topology of  $Y = \{X_\sigma\}$ , which is equivalent to the topology induced from the natural topology of the left coset space  $G/(G_0 \cdot I)$  under the mapping  $X_\sigma \rightarrow \sigma G_0 I$ . On the other hand

the factor group  $G/G_0$  is regarded as a group of isometric transformations of  $Y$ .

**Lemma 3.10.** *If  $X$  and  $G$  satisfy the conditions, i), iv), v) on page 53, then  $Y$  and  $G/G_0$  satisfy the same conditions.*

**Proof.** i, iv) are evident (see D and E on p. 60 of [20]). Let  $\sigma_1, \sigma_2, \dots$  be a sequence for which  $\sigma_1 X_\sigma, \sigma_2 X_\sigma, \dots$  forms a Cauchy sequence on  $Y$  for every  $X_\sigma$ . In the other word,  $\sigma_1, \sigma_2, \dots$  is a sequence, for which there corresponds a sequence  $\tau_1, \tau_2, \dots$  of elements of  $G_0$  such that

$$((\sigma_n \tau_n \sigma x_0, \sigma_m \tau_m \sigma x_0)) \rightarrow 0 \text{ as } n, m \rightarrow \infty$$

for every  $\sigma \in G$ . By the condition v) assumed on  $X$  and  $G$ , there corresponds  $\sigma_0 \in G$  such that

$$\lim_{n \rightarrow \infty} \sigma_n \tau_n \sigma x_0 = \sigma_0 \sigma x_0 \text{ for every } \sigma \in G.$$

This means that

$$\lim_{n \rightarrow \infty} \sigma_n X_\sigma = \sigma_0 X_\sigma \text{ for every } X_\sigma \in Y.$$

By this Lemma,  $G/G_0$  can be considered as a  $\sigma$ -compact and locally compact topological group. This topology of  $G/G_0$  is equivalent to the natural topology of the factor group of the topological group  $G$  modulo  $G_0$ , because of the equality  $\{\rho: ((\rho X_\sigma, X_\sigma)) < \varepsilon \text{ for } \sigma \in k \text{ (compact set)}\} = \{\rho: ((\rho \sigma x_0, \sigma x_0)) < \varepsilon \text{ for } \sigma \in k\} \cdot G_0$  and of the compactness of  $\{\sigma x_0: \sigma \in k\}$  and of  $\{X_\sigma: \sigma \in k\}$  in the respective spaces  $X$  and  $Y$ .

From Lemma 3.4 follows that  $G/G_0$ , a group of isometric transformations of  $Y$ , is isomorphic to  $\overline{G^Y}$ , a group of isometric transformations of  $\Omega^Y$ , as a topological group, if  $P_Y$  is absolutely continuous with respect to the invariant measure on  $Y = \{X_\sigma\}$  (see (3.4) below). Hence the homomorphism  $\sigma \rightarrow \bar{\sigma}_Y$  of  $G$  onto  $\overline{G^Y}$  is open and continuous.

Define  $m_Y(C) = m(\{\sigma: X_\sigma \in C\})$  for every Borel subset  $C$  of  $Y$ . This measure is a unique  $G/G_0$ -invariant measure on  $Y$ .

Since  $X_\sigma$  is a homogeneous space with respect to  $G_0$ , and since  $G_0$  (or more precisely  $G_0 / (\bigcap_{\tau \in G_0} \tau \sigma I \sigma^{-1} \tau^{-1})$ ) can be regarded as a group of isometric transformations of the compact space  $X_\sigma$  which satisfies the conditions i)–iv) of page 53, there is one and only one  $G_0$ -invariant measure  $n_\sigma$  of total measure 1 on each  $X_\sigma$  (see Lemmas 3.7 and 3.8). These measures fulfil evidently the relation

$$n_\sigma(B_\sigma) = n_\sigma(\sigma^{-1} B_\sigma), \quad B_\sigma = B \cap X_\sigma, \quad B \in \mathfrak{B},$$

and hence it holds that

$$(3.3) \quad m(B \cap C^*) = \int_\sigma n_\sigma(B_\sigma) m_Y(dX_\sigma),$$

where  $B_\sigma = B \cap X_\sigma$  and  $C^* = \{\sigma x_0: X_\sigma \in C\}$ . In fact, let

$$Q(B \cap C^*) = \int_C n_\sigma(B_\sigma) m_Y(dX_\sigma).$$

Then

$$\begin{aligned} Q(\rho(B \cap C^*)) &= Q(\rho B \cap (\rho C)^*) \\ &= \int_{\rho C} n_\sigma(\sigma^{-1}(\rho B)_\sigma) m_Y(dX_\sigma) \\ &= \int_{\rho C} n_\sigma(\sigma^{-1} \rho B_{(\rho^{-1}\sigma)}) m_Y(dX_\sigma) \\ &= \int_C n_\sigma(\sigma^{-1} B_\sigma) m_Y(dX_\sigma) = Q(B \cap C^*). \end{aligned}$$

From the uniqueness of the  $G$ -invariant measure on  $X$ , we have (3.3).

For an absolutely continuous measure  $P$  with respect to  $m$  on  $X$ , we have

$$\begin{aligned} P(B \cap C^*) &= \int_{B \cap C^*} p(x) m(dx) \\ &= \int_C m_Y(dX_\sigma) \int_{B_\sigma} p(x) n_\sigma(dx). \end{aligned}$$

Since  $P(X)=1$ , the integral

$$\int_{X_\sigma} p(x) n_\sigma(dx)$$

has a finite value for  $m_Y$ -almost every  $X_\sigma$ . Write

$$p_n(X_\sigma) = \int_{X_\sigma} p(x) n_\sigma(dx).$$

Then

$$P(B \cap C^*) = \int_C \frac{\int_{B_\sigma} p(x) n_\sigma(dx)}{\int_{X_\sigma} p(x) n_\sigma(dx)} p_n(X_\sigma) m_Y(dX_\sigma)$$

and

$$(3.4) \quad P_Y(C) = P(C^*) = \int_C p_n(X_\sigma) m_Y(dX_\sigma).$$

Hence we can define the conditional probability as

$$P(B : X_\sigma) = \frac{\int_{B_\sigma} p(x) n_\sigma(dx)}{\int_{X_\sigma} p(x) n_\sigma(dx)}.$$

From the definition of  $G_0$  we have

$$p(\tau x) = p(x) \text{ } m\text{-almost everywhere on } X \text{ for } \tau \in G_0.$$

Let  $S_\sigma = \{x \in X_\sigma : p(x)/p_n(X_\sigma) > 1\}$ , and  $R_\sigma = \{x \in X_\sigma : p(x)/p_n(X_\sigma) < 1\}$ . If there is a subset  $C$  of  $Y$  which is of positive  $m_Y$ -measure, and on each  $X_\sigma$  of which  $n_\sigma(S_\sigma) > 0$  and  $n_\sigma(R_\sigma) = 0$  hold, then  $P(X) > 1$ . This is impossible. Hence  $n_\sigma(S_\sigma) > 0$  implies  $n_\sigma(R_\sigma) > 0$  for  $m_Y$ -almost all  $X_\sigma$ . However

this is a contradiction, for there exists an element  $\tau \in G_0$  such that  $n_{\sigma}(\tau S_{\sigma} \cap R_{\sigma}) > 0$ , from x) of Theorem 1.4.

Thus we have

$$p(x) = p_n(X_{\sigma}) \text{ for } x \in X_{\sigma}, \text{ m-almost everywhere on } X,$$

and hence

$$P(B: X_{\sigma}) = n_{\sigma}(B_{\sigma}) \text{ m}_Y\text{-almost everywhere on } Y.$$

From this follows

**Lemma 3.11.** *A  $G_0$ -invariant statistic exists and satisfies 3), 4) and 5) of Assumption C. And*

$$(3.5) \quad P(B \cap C^*) = \int_C n_{\sigma}(B_{\sigma}) p_n(X_{\sigma}) m_Y(dX_{\sigma}).$$

Therefore, from this lemma and b) of section 1.2, we can see that the mapping  $\bar{\sigma}P \rightarrow \bar{\sigma}_Y P_Y$  is isometric and hence  $\bar{G}_Y = \{\bar{e}\}$ , and that  $\bar{G}$  is isomorphic to  $\bar{G}^Y$  as a topological group.

**Lemma 3.12.** *Suppose that  $g_0$  is a locally compact group, and*

$$g_1 \supset g_2 \supset \dots \supset g_n \supset \dots$$

*is a sequence of closed subgroups of  $g_0$  such that  $g_{n+1}$  is a normal subgroup of  $g_n$  ( $n=0, 1, 2, \dots$ ) and that*

$$(3.6) \quad \bigcap_{n=1}^{\infty} g_n = \{e\}.$$

*If factor groups  $g_n/g_{n+1}$  ( $n=0, 1, \dots$ ) are all compact, then  $g_0$  is compact.*

**Proof.** If  $g_0/g_n$  is compact, then  $g_0/g_{n+1}$  is so, since  $g_0/g_n$  is isomorphic to  $(g_0/g_{n+1})/(g_n/g_{n+1})$  and since  $g_n/g_{n+1}$  is compact. By the mathematical induction every  $g_0/g_n$  is a compact group ( $n=1, 2, \dots$ ). Denote by  $g^*$  the direct product of topological groups  $g_0/g_1, g_0/g_2, \dots$ , and by  $g'$  be a closed subgroup of  $g^*$  consisting of all elements  $(\sigma g_1, \sigma g_2, \dots, \sigma g_n, \dots)$ :  $\sigma \in g_0$ . By Tichonoff's theorem  $g^*$  and hence  $g'$  are compact. By (3.6) the correspondence  $\sigma \rightarrow (\sigma g_1, \sigma g_2, \dots, \sigma g_n, \dots)$  of  $g_0$  onto  $g'$  is 1:1, and is evidently an open isomorphism. Hence the compactness of  $g_0$  follows from that of  $g'$ .

**Lemma 3.13.** *The statement 2) of Assumption C is fulfilled.*

**Proof.** Let  $S$  be an everywhere dense countable subset of  $X$ , and write  $J_n = \{\tau: \tau x_i = x_i, i=1, 2, \dots, n\} \neq \{e\}$ . Then the subset  $S_n = \{x: J_n x = x\}$  of  $X$  is closed in  $X$ , since  $y_i \in S_n$  and  $y_i \rightarrow x''$  imply  $x'' \in S_n$ . Denote by  $x'$  one of the boundary points of  $S_n$ . Such a point exists because  $X$  is connected by the condition i). Let  $v(x', \varepsilon)$  be a neighbourhood of  $x'$  whose closure is compact, and  $x_{n+1}$  a point of  $(X - S_n) \cap v(x', \varepsilon) \cap S$ . Since every element of  $J_n$  is a isometric transformation,  $((\tau x_{n+1}, x')) = ((x_{n+1},$

$x') < \varepsilon$  for every  $\tau \in J_n$ . Hence  $J_n x_{n+1}$  is closed and contained in  $v(x', \varepsilon)$  and hence compact. By Theorem 1.2  $J_n/J_{n+1}$  is also compact.

By the repetition of the above procedure we have a sequence

$$J_1 \supset J_2 \supset \dots \supset J_n \supset \dots$$

of subgroups of  $G$ , where  $J_{n+1}$  is a normal subgroup of  $J_n$  ( $n=1, 2, \dots$ ) and

$$\bigcap_{n=1}^{\infty} J_n = \{e\}. \quad (\text{by the remark of the footnote 8}).$$

Hence the compactness of  $J_1$  follows from Lemma 3.12.

**Theorem 3.1.** *Suppose that a sample space  $X$  and a transitive group  $G$  on  $X$  satisfy the conditions i)–vi) on page 53. Then  $\Omega$  satisfies Assumptions A and B, and hence  $\bar{G}$  can be topologized. Further the homomorphism  $\sigma \rightarrow \bar{\sigma}$  of  $G$  is open and continuous, and the kernel of this homomorphism is  $\bar{G}_0$ . Assumptions C and D are fulfilled.*

**Proof.** From Lemmas 3.10, 3.11 and 3.4,  $\bar{G}^Y$  can be topologized such that the isomorphism  $\sigma G_0 \rightarrow \bar{\sigma}_Y$  of  $G/G_0$  onto  $\bar{G}^Y$  is topological. Since  $\Omega^Y$  fulfils Assumption A (by Lemmas 3.10 and 3.5) and since  $\bar{\sigma}P \rightarrow \bar{\sigma}_Y P_Y$  is an isometric mapping of  $\Omega$  onto  $\Omega^Y$  (by (3.5)),  $\Omega$  also satisfies Assumption A. Hence  $\bar{G}$  can be regarded as a topological group, where  $U(K, \varepsilon)$  forms a complete system of neighbourhoods of  $\bar{e}$ , and the mapping  $\bar{\sigma} \rightarrow \bar{\sigma}_Y$  is a topological isomorphism of  $\bar{G}$  onto  $\bar{G}^Y$  as seen in Section 1.2. Thus we can see easily that the homomorphism  $\sigma \rightarrow \bar{\sigma}$  is open and continuous.

Now we shall prove that  $(\Omega, \bar{G})$  satisfies Assumption B. By Lemma 3.6 and 3.10,  $(\Omega^Y, G^Y)$  satisfies Assumption B. Suppose that  $\sigma_{1\rho}P, \sigma_{2\rho}P, \dots$  is a Cauchy sequence in  $\Omega$  for every  $\bar{\rho}P$ . Since  $\bar{\rho}P \rightarrow \bar{\rho}_Y P_Y$  is an isometric mapping,  $(\sigma_{1\rho})_Y P_Y, (\sigma_{2\rho})_Y P_Y, \dots$  is a Cauchy sequence in  $\Omega^Y$  for every  $\bar{\rho}_Y P_Y$ . Hence there is a  $\bar{\sigma}_Y \in \bar{G}^Y$  such that

$$\lim_{n \rightarrow \infty} (\sigma_{n\rho})_Y P_Y = \bar{\sigma}_Y \bar{\rho}_Y P_Y, \quad \bar{\rho}_Y P_Y \in \Omega^Y.$$

By applying the isometric mapping  $\bar{\sigma}_Y P_Y \rightarrow \bar{\sigma}P$  again, we have

$$\lim_{n \rightarrow \infty} \overline{\sigma_{n\rho}P} = \overline{\sigma_0 \rho P}, \quad \bar{\rho}P \in \Omega,$$

where  $\sigma_0$  is an element such that  $(\sigma_0)_Y = \bar{\sigma}_Y$ .

The last statement follows directly from Lemmas 3.11 and 3.13.

**Remark 3.2.** In this section we have assumed that the sample space  $X$  is a metric space, and that  $G$  is a group of isometric transformations of  $X$ . However a  $G$ -invariant metric can be defined in  $X$ , if  $X$  is a uniform space satisfying the first axiom of countability, where the neigh-

neighbourhood  $V_n(x)$  is symmetric and  $G$ -invariant, i.e.

- a)  $y \in V_n(x)$  implies  $x \in V_n(y)$ ,
- b)  $\sigma V_n(x) = V_n(\sigma x)$  for all  $\sigma \in G$ ,
- c)  $\bigcap_{n=1}^{\infty} V_n(x) = \{x\}$ .

(see Bourbaki [5, Ch. IX, p. 22]).

### 3.2. Reductions to a homogeneous space.

**Lemma 3.14.** *Suppose that  $X$  is the direct product of two metric spaces  $Y$  and  $Z$  admitting transitive groups  $G$  and  $O$  of isometric transformations respectively. If  $X$  and the group  $G' = G \times O$  satisfy the conditions i)–v) of the beginning of the present section, then  $Y$  and  $G$  also satisfy these conditions. Further the topology of  $G$  induced by the metric of  $Y$  is equivalent to the relative topology of  $G$  as a subspace of the topological group  $G'$ .*

**Proof.** Since the conditions i)–iii) and v) are evidently satisfied by  $Y$  and  $G$ , we shall show only that iv) is fulfilled. If  $\sigma_1 y, \sigma_2 y, \dots$  is a Cauchy sequence of points for every  $y \in Y$ , then  $(\sigma_1 y, z), (\sigma_2 y, z), \dots$  is also a Cauchy sequence in  $X$  for every  $y \in Y$  and  $z \in Z$ . Hence there exists an element of  $\sigma_0 \in G$  such that

$$\lim_{n \rightarrow \infty} (\sigma_n y, z) = (\sigma_0 y, z) \text{ for every } (y, z) \in X,$$

that is,

$$\lim_{n \rightarrow \infty} \sigma_n y = \sigma_0 y \text{ for every } y \in Y.$$

The last statement of our Lemma follows directly from that  $((\sigma \tau x, \tau x)) = ((\sigma x, x))$  holds for every  $\sigma \in G, \tau \in O$  and  $x \in X$ .

**Theorem 3.2.** *Under the same assumptions as Lemma 3.14,  $\Omega$  satisfies Assumption A,  $\overline{G}$  becomes a topological group, and  $G$  is homomorphic to  $\overline{G}$  as a topological group, if the initial probability measure  $P$  defined on  $X$  and the  $\sigma$ -field  $\mathfrak{B}$  of all Borel subsets of  $X$  is absolutely continuous with respect to the  $G'$ -invariant measure on  $X$ .*

**Proof.** By Theorem 3.1. the mapping  $\sigma \rightarrow \bar{\sigma}$  of  $G'$  onto  $\overline{G}'$  is open and continuous, where  $\overline{G}'$  is a group of transformations on the metric space  $\Omega' = \{\bar{\sigma}P : \sigma \in G'\}$ . Under this mapping  $G$  is mapped onto  $\overline{G}$ , and, since  $G$  is closed in  $G'$ ,  $\overline{G}$  is closed in  $\overline{G}'$ . Hence  $\Omega$  is closed in  $\Omega'$ , and satisfies Assumption A. From the relation  $V(\sigma \tau \rho P, \tau \rho P) = V(\sigma \rho P, \rho P)$  for every  $\tau \in O$  and  $\sigma, \rho \in G$ , it follows that the topology of  $\overline{G}$  induced by the absolute variation metric  $V$  of  $\Omega$  is equivalent to the relative topology of  $\overline{G}$  as a subspace of  $\overline{G}'$ . Thus we can see that  $G$  is homomorphic to  $\overline{G}$  under the mapping  $\sigma \rightarrow \bar{\sigma}$  as a topological group.



**Theorem 3.3.** *Suppose that  $(S, \mathfrak{S}, p)$  is a probability space, where  $(S, \mathfrak{S})$  admits a group  $\Gamma$  of transformations satisfying the conditions i)–v) in the beginning of the present paper, and that  $p$  is absolutely continuous with respect to the  $\Gamma$ -invariant measure on  $S$ . Let  $(X, \mathfrak{B}, P)$  be the direct product probability space of a finite number  $n$  of  $(S, \mathfrak{S}, p)$ , and  $G$  be the group of all transformations  $(\sigma, \sigma, \dots, \sigma): (s_1, s_2, \dots, s_n) \rightarrow (\sigma s_1, \sigma s_2, \dots, \sigma s_n)$ . Then the distribution space  $\Omega = \{(\overline{\sigma, \sigma, \dots, \sigma})P; \sigma \in \Gamma\}$  satisfies Assumption A, and hence  $\overline{G}$  becomes a topological group (by Theorem 1.1), in each of the following cases:*

- (a)  $\Gamma$  is commutative,
- (b)  $S$  is compact.

*Further, in such cases the mapping  $\sigma \rightarrow (\sigma, \sigma, \dots, \sigma)$  of  $\Gamma$  onto  $\overline{G}$  is open and continuous.*

**Proof.** Denote by  $G'$  the direct product group  $\Gamma \times \Gamma \times \dots \times \Gamma$  of  $n$  groups  $\Gamma: (\sigma_1, \sigma_2, \dots, \sigma_n) \cdot (s_1, s_2, \dots, s_n) = (\sigma_1 s_1, \sigma_2 s_2, \dots, \sigma_n s_n)$ . The space  $X$  and the group  $G'$  on  $X$  satisfy the conditions i)–v) in p. 53. Hence by Mibu's theorem (Theorems 1.1 and 1.2)  $G'$  can be topologized by the metric of  $X$  such that the operation of product is continuous. Evidently  $G$  is closed in  $G'$  and is isomorphic to  $\Gamma$  as a topological group.

*Case (a).* Denote by  $O$  the subgroup of  $G'$  consisting of the elements whose first coordinate  $\sigma_1$  coincides with the neutral element  $e$  of  $\Gamma$ . Since  $\Gamma$  is commutative, and since the decomposition  $(\sigma_1, \sigma_2, \dots, \sigma_n) = (\sigma_1, \sigma_1, \dots, \sigma_1) \cdot (e, \sigma_1^{-1} \sigma_2, \dots, \sigma_1^{-1} \sigma_n)$  is unique,  $G'$  is the direct product of  $G$  and  $O$ . Moreover, fixing a point  $(s^0, \dots, s^0)$  of  $X$ , it follows from the commutativity of  $\Gamma$  that the mapping  $(\sigma_1, \sigma_2, \dots, \sigma_n) \rightarrow (\sigma_1 s^0, \dots, \sigma_n s^0)$  of  $G'$  onto  $X$  is 1:1. Hence  $X$  may be considered as a direct product space  $Y \times Z: Y = \{(s, s, \dots, s): s \in S\}$  and  $Z = \{(s^0, \sigma_2 s^0, \dots, \sigma_n s^0): \sigma_i \in \Gamma, i=2, \dots, n\}$ , and  $G$  and  $O$  are transitive groups of transformations on  $Y$  and  $Z$  respectively. Therefore from Theorem 3.2 follows the conclusion.

*Case (b).* By Theorem 3.1  $\overline{G}'$ , a group of transformations of  $\Omega' = \{\overline{\sigma}P: \sigma \in G'\}$  induced by  $G'$ , becomes a topological group and the mapping  $(\sigma_1, \sigma_2, \dots, \sigma_n) \rightarrow (\overline{\sigma_1}, \overline{\sigma_2}, \dots, \overline{\sigma_n})$  of  $G'$  onto  $\overline{G}'$  is a homomorphism topologically. Hence  $\sigma \rightarrow (\overline{\sigma, \sigma, \dots, \sigma})$  of  $\Gamma$  onto  $\overline{G}$  as a subspace of  $\overline{G}'$ , is open and continuous. Since  $G$  is closed in  $G'$ ,  $\overline{G}$  is closed in  $\overline{G}'$ . Hence  $\Omega$  is closed in  $\Omega'$ , and satisfies Assumption A. Hence  $\overline{G}$  becomes a topological group as a group of transformations of  $\Omega$ , by the aid of the absolute variation metric of  $\Omega$ . On the other hand, the identical mapping of the subgroup  $\overline{G}$  of  $\overline{G}'$  onto the group  $\overline{G}$  of transformations of  $\Omega$  is continuous, because every neighbourhood  $U[\{(\overline{\sigma, \sigma, \dots, \sigma})P\}, \varepsilon]$ , in  $\overline{G}$ , of  $\bar{e}$  is the intersection of  $\overline{G}$  and a neighbourhood, in  $\overline{G}'$ , of  $\bar{e}$ . From the compactness of  $S$  follows that of  $I'$ , of  $G$ , and of the sub-

group  $\bar{G}$  of  $\bar{G}'$ . Therefore the mapping  $\sigma \rightarrow (\sigma, \sigma, \dots, \sigma)$  is open and continuous.

#### 4. Some examples.

a) *The permutation parameter.* Let  $S$  be a set of  $k$  mutually exclusive events  $s_1, s_2, \dots, s_k$  with probability  $p_1, p_2, \dots, p_k$  ( $\sum p_i = 1$ ) respectively, and denote such a probability space by  $(S, p)$ . Consider the group  $\Gamma$  of the permutations  $\sigma$  of  $1, 2, \dots, k$ . Every element  $\sigma$  of  $\Gamma$  introduces an another probability space  $(S, \bar{\sigma}p)$  as follows:

$$\bar{\sigma}p(s_i) = p(s_{\sigma^{-1}(i)}) = p_{\sigma^{-1}(i)}; i = 1, 2, \dots, k.$$

Suppose that  $(X, \bar{\sigma}P)$  is a direct product probability space  $(S \times S \times \dots \times S, \bar{\sigma}p \times \bar{\sigma}p \times \dots \times \bar{\sigma}p)$  of  $n$  same probability spaces  $(S, \bar{\sigma}p)$ , *i.e.* that  $X$  is a set of the sample points  $x = (x_1, \dots, x_n)$  in  $n$  independent trials, where the probability of occurrence of  $n_1$  events of the first kind,  $n_2$  of the second kind, ... etc. ( $\sum n_i = n$ ) in some order is give as

$$(4.1) \quad p_{\sigma^{-1}(1)}^{n_1} p_{\sigma^{-1}(2)}^{n_2} \dots p_{\sigma^{-1}(k)}^{n_k}.^{12)}$$

Let  $G'$  be the group of the transformations  $\sigma: (x_1, \dots, x_n) \rightarrow (\sigma x_1, \sigma x_2, \dots, \sigma x_n)$  of  $X$  for  $\sigma \in \Gamma$ , and  $G_0$  the group of the permutations of the coordinates of  $x = (x_1, x_2, \dots, x_n)$ . Evidently every element of  $G_0$  leaves (4.1) invariant, and the group  $G = G' \times G_0$  of transformations on  $X$  induces a distribution space  $\Omega = \{\bar{\sigma}P: \sigma \in G\}$ . Hence the  $G_0$ -invariant statistic  $y(x) = (n_1, n_2, \dots, n_k)$  for the above  $x$  is sufficient for  $\Omega$ , and the distribution  $\bar{\sigma}_Y P_Y$  on the range  $Y$  of  $y(x)$  is given as

$$\bar{\sigma}_Y P_Y[(n_1, n_2, \dots, n_k)] = \frac{n!}{n_1! n_2! \dots n_k!} p_{\sigma^{-1}(1)}^{n_1} p_{\sigma^{-1}(2)}^{n_2} \dots p_{\sigma^{-1}(k)}^{n_k},$$

when  $\bar{\sigma}P$  is the true distribution of  $x$ .

If  $(n_1, n_2, \dots, n_k)$  is a permutation of a non-decending sequence  $(m_1, m_2, \dots, m_k)$ , we put

$$z(n_1, n_2, \dots, n_k) = (m_1, m_2, \dots, m_k)$$

We can see easily that  $z(y(x))$  is a  $G$ -invariant statistic, and that its distribution is

$$P_Z(m_1, m_2, \dots, m_k) = \frac{n!}{m_1! m_2! \dots m_k!} \sum p_1^{m_1} p_2^{m_2} \dots p_k^{m_k}$$

irrespectively of the distribution of  $x$ , where the summation is over all different permutations  $(n_1, n_2, \dots, n_k)$  of  $(m_1, m_2, \dots, m_k)$ . Hence the conditional probability of  $(n_1, n_2, \dots, n_k)$  for given  $z(y) = (m_1, m_2, \dots, m_k)$  when  $\sigma P$  is true is

<sup>12</sup> (12) For the simplicity we shall assume that (4.1) has a distinct value for a distinct set  $(n_1, n_2, \dots, n_k)$ .

$$\frac{p_{\sigma^{-1}(1)}^{n_1} p_{\sigma^{-1}(2)}^{n_2} \dots p_{\sigma^{-1}(k)}^{n_k}}{\sum p_1^{n_1} p_2^{n_2} \dots p_k^{n_k}}$$

Evidently  $\bar{G}$  is a finite group, and hence compact. The fiducial measure  $\bar{P}(\bar{\sigma}:z)$  is such that

$$\bar{P}(\bar{\sigma}:z) = \frac{p_1^{m_{\sigma(1)}} p_2^{m_{\sigma(2)}} \dots p_k^{m_{\sigma(k)}}}{\sum_{\sigma \in I^k} p_1^{m_{\sigma(1)}} p_2^{m_{\sigma(2)}} \dots p_k^{m_{\sigma(k)}}$$

Suppose for example that  $\mathbf{A} = \mathcal{Q} = \{\bar{\sigma}P\}$  and  $W(\bar{\sigma}P, \bar{\tau}P) = \sum_{i=1}^k (p_{\sigma^{-1}(i)} - p_{\tau^{-1}(i)})^2$ . Then we have

$$\begin{aligned} \mathbf{b}(\bar{\sigma}P, z) &= \sum_{\rho \in I^k} \sum_{i=1}^k (p_i - p_{\sigma^{-1}\rho^{-1}(i)})^2 \bar{P}(\bar{\rho}:z) \\ &= \sum_{\rho \in I^k} \sum_{i=1}^k (p_i^2 + p_{\sigma^{-1}\rho^{-1}(i)}^2 - 2p_i p_{\sigma^{-1}\rho^{-1}(i)}) \bar{P}(\bar{\rho}:z) \\ &= 2 \sum_{\rho \in I^k} \sum_{i=1}^k (p_i^2 - p_i p_{\sigma^{-1}\rho^{-1}(i)}) \bar{P}(\bar{\rho}:z) \\ &= 2 \left[ \left( \sum_{i=1}^k p_i^2 \right) - \left( \sum_{i=1}^k p_i \sum_{\rho \in I^k} p_{\sigma^{-1}\rho^{-1}(i)} \bar{P}(\bar{\rho}:z) \right) \right] \\ &= 2 \left[ \sum_{i=1}^k p_i^2 - b'(\sigma) \right], \text{ (say).} \end{aligned}$$

Hence the PGSB d.f. is such that if the observed number  $(n_1, n_2, \dots, n_k)$  is a result by permutation  $\sigma$  of  $z = (m_1, m_2, \dots, m_k)$ , we decide with probability 1 that  $\bar{\sigma}\bar{\tau}_z P$  is the true distribution, where  $\tau_z$  maximizes

$$b'(\tau) = \sum_{i=1}^k p_i \sum_{\rho \in I^k} p_{\tau^{-1}\rho^{-1}(i)} \bar{P}(\bar{\rho}:z)$$

For  $k=2$ ,

$$b'(\sigma_0) M = p_1^{m_1} p_2^{m_2} (p_1^2 + p_2^2) + 2p_1^{m_2+1} p_2^{m_1+1}$$

for the identical permutation  $\sigma_0$ , and

$$b'(\sigma_1) M = 2p_1^{m_1+1} p_2^{m_2+1} + (p_1^2 + p_2^2) p_1^{m_2} p_2^{m_1}$$

for  $\sigma_1 = (12)$ , where  $M = p_1^{m_1} p_2^{m_2} + p_1^{m_2} p_2^{m_1}$ . Hence if  $p_1 < p_2$ , then

$$p_1^{m_1} p_2^{m_2} \geq p_1^{m_2} p_2^{m_1}$$

and we have  $b'(\sigma_0) \geq b'(\sigma_1)$ . This shows that the estimating process, which indicates  $\sigma_1 P$  or  $P$  according as the number  $n_1$  of occurrences of the event  $s_1$  is greater or smaller than  $n_2$  of  $s_2$ , is a PGSB d.f.

For  $k=3$ , suppose that  $p_1 < p_2 < p_3$ , and denote

$$\begin{aligned} a_0 &= p_1^2 + p_2^2 + p_3^2, & a_1 &= p_1^2 + 2p_2 p_3, & a_2 &= p_2^2 + 2p_1 p_3, \\ a_3 &= p_3^2 + 2p_1 p_2, & a_4 &= p_1 p_2 + p_2 p_3 + p_3 p_1, \end{aligned}$$

and

$$P_0 = p_1^{m_1} p_2^{m_2} p_3^{m_3}, \quad P_1 = p_1^{m_1} p_2^{m_3} p_3^{m_2}, \quad P_2 = p_1^{m_2} p_2^{m_2} p_3^{m_1},$$

$$P_3 = p_1^{m_2} p_2^{m_1} p_3^{m_3}, \quad P_4 = p_1^{m_2} p_2^{m_3} p_3^{m_1}, \quad P_5 = p_1^{m_3} p_2^{m_1} p_3^{m_2}.$$

Then we have

$$(4.2) \quad a_0 > a_1 > a_4 > a_2, \quad a_0 > a_3 > a_4 > a_2,$$

$$P_0 > P_1 > P_4 > P_2, \quad P_1 > P_5,$$

$$(4.3) \quad P_0 > P_3 > P_5 > P_2, \quad P_3 > P_4.$$

Since the group  $G'$  consists of elements  $\sigma_0 = \text{identical}$ ,  $\sigma_1 = (12)$ ,  $\sigma_2 = (23)$ ,  $\sigma_3 = (31)$ ,  $\sigma_4 = (132)$  and  $\sigma_5 = (123)$ , we have

$$b'(\sigma_0) = a_0 P_0 + a_1 P_1 + a_3 P_3 + a_4 P_4 + a_4 P_5 + a_2 P_2,$$

$$b'(\sigma_1) = a_3 P_0 + a_4 P_1 + a_0 P_3 + a_2 P_4 + a_1 P_5 + a_4 P_2,$$

$$b'(\sigma_2) = a_1 P_0 + a_0 P_1 + a_4 P_3 + a_3 P_4 + a_2 P_5 + a_4 P_2,$$

$$b'(\sigma_3) = a_2 P_0 + a_4 P_1 + a_4 P_3 + a_1 P_4 + a_3 P_5 + a_0 P_2,$$

$$b'(\sigma_4) = a_4 P_0 + a_3 P_1 + a_2 P_3 + a_0 P_4 + a_4 P_5 + a_1 P_2,$$

$$b'(\sigma_5) = a_4 P_0 + a_2 P_1 + a_1 P_3 + a_4 P_4 + a_0 P_5 + a_3 P_2,$$

Therefore, following from (4.2) and (4.3)

$$b'(\sigma_0) = \max_{0 \leq i \leq 5} b'(\sigma_i).$$

This shows that the PGSB d.f. is a non-randomized estimation such that if the triplet  $(n_1, n_2, n_3)$  of the numbers of occurrences of the events  $s_1, s_2, s_3$  is a permutation  $\sigma$  of  $(m_1, m_2, m_3)$ , then the distribution  $\bar{\sigma}P$  is accepted.

For  $k \geq 4$ , calculations are rather troublesome.

b) *The discrete location parameter.* Let  $X$  be Euclidean  $n$ -space  $R^n$ , and  $P$  is a discrete distribution which distributes on  $x_{ij} = y_i + r_{ij}x_0$  the probability  $s_i p_{ij}$ , where  $y_i = (y_{i1}, y_{i2}, \dots, y_{in})$  is a point in the hyperplane  $\sum_{i=1}^n y_{ii} = 0$ ,  $x_0 = (1, 1, \dots, 1)$ ,  $s_i > 0$ ,  $\sum_{i=1}^n s_i = 1$ ,  $p_{ij} \geq 0$ ,  $\sum_{j=1}^n p_{ij} = 1$ , and  $i, j = 1, 2, \dots$ .

The group  $G$  of translations  $h: x \rightarrow x + hx_0$  of  $X$  induces the distribution space  $\Omega = \{\bar{h}P\}$ :

$$\bar{h}P(x_{ij} + hx_0) = s_i p_{ij}.$$

However  $\Omega$  is discrete and non-countable, in fact

$$V(\bar{h}P, P) \geq \sup_{\substack{1 \leq i < \infty \\ 1 \leq j < \infty}} \inf_{\substack{1 \leq i' < \infty \\ 1 \leq j' < \infty}} \max [ |s_i p_{ij} - s_{i'} p_{i'j'}|, s_i p_{ij} ] > 0,$$

whatever  $h$  may be. Hence  $\Omega$  is not necessarily  $\sigma$ -compact, so that Mibu's Theorems 1.1-1.4 can not be directly applied for this problem.

Let  $h_1, h_2, \dots$  be a sequence of real numbers such that

1)  $n_1 h_1 + n_2 h_2 + \dots + n_k h_k = 0$  implies  $n_1 = n_2 = \dots = n_k = 0$  for every integer  $k \geq 1$ ,

2) for every  $r_{ij}$  there is one and only one linear expression

$$r_{ij} = \sum_{\alpha=1}^k n_{\alpha} h_{\alpha}$$

with integral coefficients  $n_1, n_2, \dots, n_k$ .

Denote by  $G'$  the normal subgroup of  $G$  consisting of all linear combination  $\sum n_{\alpha} h_{\alpha}$  of every finite number of  $h_1, h_2, \dots, h_k$  with integral coefficients  $n_1, n_2, \dots, n_k$ . The distribution space  $\Omega'$ , induced from  $P$  by  $G'$ , is a discrete but countable space, and hence is  $\sigma$ -compact. Evidently  $\bar{G}$  may be considered to coincide with  $G'$ . Let  $J_k$  be the subset of  $G'$  consisting of all  $\sum_{\alpha=1}^k n_{\alpha} h_{\alpha}$  with  $|n_{\alpha}| < k (\alpha=1, 2, \dots, k)$  for a fixed positive integer  $k$ . Then we have

$$\bigcup_{k=0}^{\infty} J_k = G'$$

and  $J_k$  is a finite set. Hence there is an integer  $k > 0$  for every finite subset  $R$  such that  $J_k \supset R$ . It is obvious that  $G'$  is an  $A$ -group. Hence, by Theorem 2.2, we have

$$\sup_{\bar{h}P \in \Omega} r(\bar{h}P, \varphi) \geq \sup_{\bar{h}P \in \Omega'} r(\bar{h}P, \varphi) \geq \sum_{i=1}^{\infty} F_i \text{ for every d.f. } \varphi$$

if there is a finite subset  $R$  of  $G'$  for every  $\varepsilon > 0$  and every integer  $i$  ( $0 < i < \infty$ ) such that

$$\sum_{r_{ij} \in R} w(h + r_{ij}) p_{ij} > F_i - \varepsilon \text{ for every } h \in G',$$

where  $F_i = \inf_{h \in G'} \sum_{j=1}^{\infty} w(h + r_{ij}) p_{ij}$ . This is Theorem 11.3.1. of Blackwell-Girshick's Book [4].

The admissibility of this d.f. has shown by Blackwell [2] under a preferably weak condition.

c) *The continuous location parameter.* Let  $X = R^n (n \geq 2)$ ,  $\mathfrak{B}$  the class of all Borel sets of  $R^n$ , and  $P$  a probability measure with a density function  $f(x) = f(x_1, x_2, \dots, x_n)$ :

$$P(B) = \int_B f(x_1, x_2, \dots, x_n) dx_1 dx_2 \dots dx_n, \quad B \in \mathfrak{B},$$

where the integration is the usual Lebesgue's one.

Suppose that  $G$  is a group of transformations  $(x_1, x_2, \dots, x_n) \rightarrow (x_1 + \sigma, x_2 + \sigma, \dots, x_n + \sigma)$ , where  $\sigma$  is a real number. This group  $G$  induces a distribution space  $\Omega = \{\bar{\sigma}P : -\infty < \sigma < \infty\}$  from  $P$  as follows:

$$\begin{aligned} \bar{\sigma}P(B) &= \int_{B-\sigma} f(x_1, x_2, \dots, x_n) dx_1 dx_2 \dots dx_n \\ &= \int_B f(x_1 - \sigma, x_2 - \sigma, \dots, x_n - \sigma) dx_1 dx_2 \dots dx_n. \end{aligned}$$

Such a parameter  $\sigma$  is called a *location parameter*.

By Theorem 3.2 the mapping  $\sigma \rightarrow \bar{\sigma}$  of  $G$  onto the parameter group  $\bar{G}$  is isomorphic topologically, that is,  $\bar{G}$  is also considered as the additive group of the real numbers. Hence  $\bar{G}$  is an  $A$ -group.

The fiducial measure  $\bar{P}(L:z)$  for  $z=(0, z_2, \dots, z_n)$  is

$$\bar{P}(L:z) = \int_L f(\sigma, z_2 + \sigma, \dots, z_n + \sigma) d\sigma / h(z) dz.$$

Suppose that  $\mathbf{A} = \Omega$  and  $W(\bar{\sigma}P, \bar{\tau}P) = w(\tau - \sigma)$ , and write

$$\mathbf{b}_m(\tau, z) = \int_{-m}^{+m} w(\sigma + \tau) f(\sigma, z_2 + \sigma, \dots, z_n + \sigma) d\sigma,$$

$$\mathbf{b}(\tau, z) = \int_{-\infty}^{+\infty} w(\sigma + \tau) f(\sigma, z_2 + \sigma, \dots, z_n + \sigma) d\sigma,$$

where  $z=(0, z_2, \dots, z_n)$ . If there is, for every  $z$  and every positive  $\varepsilon$ , an integer  $m$  such that

$$\mathbf{b}_m(\tau, z) > \inf_{-\infty < \tau < +\infty} \mathbf{b}(\tau, z) - \varepsilon \quad \text{for every real } \tau,$$

and if  $\tau_z$  minimizes  $\mathbf{b}(\tau, z)$ , then the PGSB d.f. exists and is the procedure in which, when  $x=(x_1, x_2, \dots, x_n)$  is observed,  $(x_1 + \tau_z)P$  is accepted, where  $z=(0, x_2 - x_1, \dots, x_n - x_1)$ .

i) If  $w(\sigma) = \sigma^2$  and if  $\int_{-\infty}^{+\infty} (\sigma + \tau)^2 f(\sigma, z_2 + \sigma, \dots, z_n + \sigma) d\sigma < \infty$  for all  $\tau$ , then we have

$$\tau_z = - \int_{-\infty}^{+\infty} \sigma f(\sigma, z_2 + \sigma, \dots, z_n + \sigma) d\sigma, \quad z=(0, z_2, \dots, z_n).$$

The minimax risk of this decision function is given as

$$\int_{-\infty}^{+\infty} \dots \int_{-\infty}^{+\infty} x_1^2 f(x_1, x_2 + x_1, \dots, x_n + x_1) dx_1 dx_2 \dots dx_n \\ - \int_{-\infty}^{+\infty} \dots \int_{-\infty}^{+\infty} \tau_z^2 h(z) dz_2 \dots dz_n.$$

ii) If  $w(\tau) = |\tau|$ , then  $\tau_z$  is such that

$$\int_{-\infty}^{\tau_z} f(\sigma, z_2 + \sigma, \dots, z_n + \sigma) d\sigma = \frac{1}{2}, \quad z=(0, z_2, \dots, z_n)$$

The above two estimates are the ones obtained by Pitman [18].

iii) If  $f(\sigma, z_2 + \sigma, \dots, z_n + \sigma)$  is a monotone decreasing function  $g_z(t)$  of  $t = |\sigma - \sigma_z|$  for every fixed  $z=(0, z_2, \dots, z_n)$  and  $w(\tau)$  is also a monotone increasing function of  $|\tau|$ , then the observed value  $(x_1, x_2, \dots, x_n)$  indicates the estimated value  $x_1 - \sigma_z$  for  $z=(0, x_2 - x_1, \dots, x_n - x_1)$ . This is a maximal likelihood estimate as well as a PGSB estimate.

iv) Suppose that  $f(x_1, x_2, \dots, x_n)$  satisfies the same conditions as the case iii),  $\mathbf{A}$  is the class of all interval  $I_{ab} = (a, b)$ :  $-\infty < a < b < +\infty$ , and that  $w(I_{ab}) = |b - a|$  if  $a < 0 < b$ ;  $= 1 + |b - a|$  if otherwise. For such

problems, the observed value  $(x_1, x_2, \dots, x_n)$  indicates the estimation by the interval  $(\sigma_z - \sigma_1 + x_1, \sigma_z + \sigma_1 + x_1)$  for  $z = (0, x_2 - x_1, \dots, x_n - x_1)$ , where  $g_z(t) \geq 1$  for  $|t| \leq \sigma_1$  and  $g_z(t) < 1$  for  $|t| > \sigma_1$ . This procedure is minimax by Theorem 2.7.

d) *The location-scale parameter.* Let  $X, \mathfrak{B}$  and  $P$  are the same as defined in c). Suppose that  $G$  is a group of transformations  $\sigma(x_1, x_2, \dots, x_n) = (\sigma x_1, \sigma x_2, \dots, \sigma x_n)$ , where  $\sigma x_i = a_\sigma x_i + b_\sigma$ ,  $+\infty > a_\sigma > 0$ ,  $+\infty > b_\sigma > -\infty$ . ( $a_{\sigma\tau} = a_\sigma a_\tau$  and  $b_{\sigma\tau} = a_\sigma b_\tau + b_\sigma$ ). Then

$$\begin{aligned} \bar{\sigma}P(B) &= P(\sigma^{-1}B) = \int_{\sigma^{-1}B} f(x_1, x_2, \dots, x_n) dx_1 dx_2 \dots dx_n \\ &= \frac{1}{a_\sigma^n} \int_B f\left(\frac{x_1 - b_\sigma}{a_\sigma}, \frac{x_2 - b_\sigma}{a_\sigma}, \dots, \frac{x_n - b_\sigma}{a_\sigma}\right) dx_1 dx_2 \dots dx_n. \end{aligned}$$

We shall consider as  $\Omega$  the whole of  $\bar{\sigma}P$ 's. In this case the pair  $(a_\sigma, b_\sigma)$  is called a *location-scale parameter*.

Write, for every point  $x = (x_1, x_2, \dots, x_n)$ ,

$$\begin{aligned} \bar{x} = \bar{x}(x) &= \frac{\sum x_i}{n}, \quad s = s(x) = \sqrt{\frac{\sum (x_i - \bar{x})^2}{n}} = \sqrt{\frac{\sum x_i^2}{n} - \bar{x}^2} \\ \text{and } z = z(x) &= \left(\frac{x_1 - \bar{x}}{s}, \frac{x_2 - \bar{x}}{s}, \dots, \frac{x_n - \bar{x}}{s}\right). \end{aligned}$$

Denote by  $X_0$  the line  $x_1 = x_2 = \dots = x_n$  and by  $x_0$  a point  $(1, 1, \dots, 1)$ . Since every  $x \notin X_0$  is expressed uniquely as

$$x = sx + \bar{x}x_0,$$

and since the  $G$ -invariant statistic  $z(x)$  has its values on the unit sphere  $Z$  of the hyperplane  $z_1 + z_2 + \dots + z_n = 0$ , we can write

$$f(x) = p(\bar{x}(x), s(x), z(x))$$

and

$$dx_1 dx_2 \dots dx_n = s^{n-2} d\bar{x} ds dz,$$

where  $dz$  is an element of area on the unit  $(n-2)$ -dimensional sphere. Denoting by  $B'$  the image of  $B (\in \mathfrak{B})$  under the mapping  $x \rightarrow (\bar{x}(x), s(x), z(x))$ , we can easily see that

$$\bar{\sigma}P(B) = \frac{1}{a_\sigma^n} \int_{B'} p\left(\frac{\bar{x} - b_\sigma}{a}, \frac{s}{a_\sigma}, z\right) s^{n-2} d\bar{x} ds dz.$$

Hence, by writing

$$h(z) = \int_0^{+\infty} \left\{ \int_{-\infty}^{+\infty} p(\bar{x}, s, z) d\bar{x} \right\} s^{n-2} ds,$$

the distribution on  $Z$  induced by the statistic  $z(x)$  is

$$\bar{\sigma}P_Z(D) = P_Z(D) = \int_D h(z) dz$$

and the conditional distribution for given  $z$  is

$$\bar{\sigma}P(B:z) = \frac{1}{a_\sigma^n} \int_{B' \cap X'_z} \frac{p\left(\frac{\bar{x}-b_\sigma}{a_\sigma}, \frac{s}{a_\sigma}, z\right)}{h(z)} s^{n-2} d\bar{x} ds,$$

where  $D$  is a Borel subset of  $Z$  and  $X'_z = \{(\bar{x}, s, z) : -\infty < \bar{x} < +\infty, 0 < s < +\infty\}$ . Hence the fiducial measure for  $z$  is such that

$$\bar{P}(L:z) = \frac{1}{h(z)} \int_L p(\bar{x}, s, z) s^{n-2} d\bar{x} ds.$$

Now consider the group  $O$  of the rotations around the fixed axis  $X_0$ . Since  $Z$  is a homogeneous space with respect to  $O$ , the algebraically homomorphic mapping  $\sigma \rightarrow \bar{\sigma}$  of  $G$  onto  $\bar{G}$  is open and continuous from Theorem 3.2. However no element  $\rho$  of  $G$  remains each  $\bar{\sigma}P$  invariant, if  $\rho$  is not the natural element, by Lemma 3.7. Hence  $\sigma \rightarrow \bar{\sigma}$  is an isomorphism as a mapping of topological groups.

Denote by  $\Omega'$  the distribution space induced from  $P$  by the direct product  $G' = G \times O$ , and by  $G_0$  the subgroup of  $G'$  remaining every distribution of  $\Omega'$  invariant. Since the intersection of  $G_0$  and  $G$  contains only the natural element as seen just now,  $G_0$  must be a normal subgroup of  $O$ . But  $O$  has no closed normal subgroup except for  $O$  itself and  $\{e\}$ . Hence  $G_0 = O$  or  $G_0 = \{e\}$ .

If  $G_0 = O$ , then  $\Omega'$  and  $\bar{G}'$  coincide with  $\Omega$  and  $\bar{G}$ , respectively. And  $(\bar{x}(x), s(x))$  becomes a sufficient statistic for  $\Omega$ .

i) Suppose that  $\mathbf{A} = \bar{G}$ , and that the loss function  $W(P, \bar{\sigma}) = w(a_\sigma, b_\sigma)$  is either a bounded function on the upper half plane  $\bar{G}$  or a continuous function satisfying

$$\lim_{a \rightarrow 0} w(a, b) = \lim_{a \rightarrow \infty} w(a, b) = \infty \quad \text{for any fixed } b,$$

and

$$\lim_{b \rightarrow \pm\infty} w(a, b) = \infty \quad \text{for any fixed } a > 0.$$

Let  $(a_z, b_z)$  be a point of  $\bar{G}$  which minimizes

$$\mathbf{b}[(a, b), z] = \frac{1}{h(z)} \int_0^\infty \left[ \int_{-\infty}^\infty w(as, sb + \bar{x}) p(\bar{x}, s, z) d\bar{x} \right] s^{n-2} ds.$$

Then the decision function  $\varphi^0$ :

$$\varphi^0[(sa_z, sb_z + \bar{x}) : sz + \bar{x}x_0] = 1 \quad \text{for every } s > 0 \text{ and } \bar{x}$$

is a PGSB d.f.

ii) Let  $g_b$  is a normal subgroup  $\{(1, b) : -\infty < b < \infty\}$ . If  $\mathbf{A} = G/g_b$ ,  $W(P, g_b(a, 0)) = w(a)$  is either a bounded function or a continuous function such that  $\lim_{a \rightarrow 0} w(a) = \lim_{a \rightarrow \infty} w(a) = \infty$ , then the invariant d.f.  $\varphi^0$ :

$$\varphi^0[(a_z s, 0) \cdot g_b : sz + x_0] = 1$$



is a PGSB d.f., where  $a_z$  minimizes

$$b(a, z) = \frac{1}{h(z)} \int_0^{+\infty} w(as) \left[ \int_{-\infty}^{+\infty} p(\bar{x}, s, z) d\bar{x} \right] s^{n-2} ds.$$

The distribution of decision under  $\varphi^0$  when  $\bar{\sigma}P$  is true is independent of  $b_\sigma$ . (See Theorem 2.6)

iii) Let  $g_a$  is a subgroup  $\{(a, 0): 0 < a < \infty\}$ . If  $\mathbf{A}$  is a right coset space of  $G$  modulo  $g_a$ , and if  $W(P, g_a \cdot (1, b)) = w(b)$  is either a bounded function or a continuous function such that  $\lim_{b \rightarrow \pm\infty} w(b) = \infty$ , then  $\varphi^0$ :

$$\varphi^0 \left( g_a \cdot \left( \frac{\bar{x} + b_z}{s} \right); sz + \bar{x}x_0 \right) = 1,$$

where  $b_z$  minimizes

$$b(b, z) = \int_{-\infty}^{+\infty} \int_0^{+\infty} w \left( \frac{\bar{x} + b}{s} \right) p(\bar{x}, s, z) s^{n-2} ds d\bar{x}.$$

The distribution of decision under this  $\varphi^0$  when  $\bar{\sigma}P$  is true depends only on  $\frac{b_\sigma + b_z}{a_\sigma}$ , if  $b_z = b_0$  for every  $z \in Z$ . (See Theorem 2.6)

If  $P$  is a normal distribution, such a decision function is of Student type.

e) *The rotation parameter.* Given  $n$  independent and identically distributed random points  $s_1, s_2, \dots, s_n$  on the unit sphere  $S$  in  $R^3$ , whose distributions are absolutely continuous with respect to the element of area on  $S$  and has a density function  $f(s)$ , we consider the space  $X$  of such samples  $x = (s_1, s_2, \dots, s_n)$  as a direct product space  $S \times S \times \dots \times S$  of  $n$  spheres. Denote by  $\sigma, \tau, \dots$  the rotations on  $S$ . The group  $G$  of transformations on  $X$ , whose elements are  $(s_1, s_2, \dots, s_n) \rightarrow (\sigma s_1, \sigma s_2, \dots, \sigma s_n)$ , induces a distribution space  $\Omega = \{\bar{\sigma}P\}$ :

$$\bar{\sigma}P(B) = P(\sigma^{-1}B) = \int_B \prod_{i=1}^n f(\sigma^{-1}s_i) ds_1 \dots ds_n.$$

By  $M(x)$  we denote the  $n \times n$  matrix  $(m_{ij}(x))$  whose  $i, j$  element  $m_{ij}(x)$  is the smaller angle between two segments joining the origin with  $s_i$  and  $s_j$  respectively. This statistic  $M(x)$  is  $G$ -invariant, and its distribution does not depend on  $\sigma$ .

Suppose that  $n \geq 3$ . By Theorem 3.3  $G$  is isomorphic to  $\bar{G}$ . Thus  $\bar{G}$  is compact and hence A-group. Therefore for this problem the PGSB d.f. exists if the loss function is continuous.

f) *The general linear transformation parameter.* Let  $X$  be the set of all  $n \times k$  matrices  $x = (x_{ij})$  ( $n > k$ ), and  $P$  be a  $k$ -variate normal distribution:

$$P(dx) = \frac{1}{(2\pi)^{\frac{1}{2}nk}} e^{-\frac{1}{2}tr(x'x)} \prod_{i,j} dx_{ij}.^{(13)}$$

(13)  $x'$  is the transposed matrix of  $x$ .

For this example, A.T. James' paper [8] may be referred. Since every regular  $k \times k$  matrix  $\sigma = (\sigma_{ij})$  defines a right transformation  $x \rightarrow x\sigma$  of  $X$  onto itself, the group  $G$  consisting of all of these transformations induces the distribution space  $\Omega = \{\bar{\sigma}P\}$ :

$$\bar{\sigma}P(dx) = \frac{|\Sigma|^{-\frac{n}{2}}}{(2\pi)^{\frac{1}{2}nk}} e^{-\frac{1}{2}tr(\Sigma^{-1}x'x)} \prod_{i,j} dx_{ij}.$$

where  $\Sigma = \sigma'\sigma$ . That is to say,  $\Omega$  consists of the sample distributions of the normal  $k$ -variates  $x_{i1}, \dots, x_{ik}$  with means zero.

$X$  is not a homogeneous space with respect to  $G$ . However, if we introduce on this space  $X$  the group  $O$  of the left transformations  $x \rightarrow \theta x$  by  $n \times n$  orthogonal matrices  $\theta$ , then  $X$  becomes a homogeneous space with respect to the direct product group  $G' = G \times O$ . Since  $X$  satisfies the conditions a), b) and c) of Remark 3.2, and iv), v) of the beginning of Section 3, and since the  $G'$ -invariant measure on  $X$  is

$$\frac{\prod_{i,j} dx_{ij}}{|x'x|^{\frac{n}{2}}},$$

the mapping  $\sigma \rightarrow \bar{\sigma}$  of  $G$  onto  $\bar{G}$  is a homomorphism, by Theorem 3.2. Evidently there is no right transformation (if not neutral) on  $X$  remaining every quadratic form  $tr(\Sigma^{-1}x'x)$  invariant. This shows that  $G_0$  consists only of the neutral element, and that  $\sigma \rightarrow \bar{\sigma}$  is an isomorphism topologically.

Thus we see that  $\bar{G}$  isomorphic to the general linear group  $G$  of the regular  $k \times k$  matrices. But we cannot yet to prove that  $G$  is an A-group. Hence we never know whether or not our method of constructing the invariant minimax decision function is valid for this problem.

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